

# Learning Parameters in Discrete Naive Bayes models by Computing Fibers of the Parametrization Map

Vincent Auvray and Louis Wehenkel

EE & CS Dept. and GIGA-R, University of Liège, Belgium

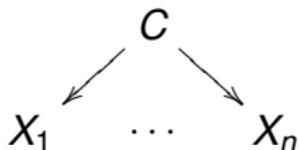
NIPS - AML 2008 - Whistler

## Naive Bayesian networks

A discrete naive Bayesian network (or latent class model) with  $m$  classes is a distribution  $p$  over discrete variables  $X_1, \dots, X_n$  such that

$$p(X_1 = x_1, \dots, X_n = x_n) = \sum_{t=1}^m p(C = t) \prod_{i=1}^n p(X_i = x_i | C = t).$$

Graphically, the independencies between  $C, X_1, \dots, X_n$  are encoded by



## Problem statement

Given a naive Bayesian network  $p$ , compute the parameters  $p(C = t)$ ,  $p(X_i = x_i | C = t)$  for  $i = 1, \dots, n$ ,  $t = 1, \dots, m$ , and  $x_i \in \mathcal{X}_i$  mapped to  $p$ .

Why?

- better understanding of the model
- estimation of parameters
- model selection
- study of parameter identifiability

# Outline

Mathematical Results

Applications

Extension

## Some notation

Given parameters of a naive Bayesian distribution, we define new parameters

$$w_t = p(C = t),$$

$$A_{x_i}^t = p(X_i = x_i | C = t) - p(X_i = x_i).$$

Given a distribution  $p$ , let

$$\begin{aligned} q(x_{i_1}, \dots, x_{i_k}) &= p(x_{i_1}, \dots, x_{i_k}) \\ &\quad - \sum_{\{X_{j_1}, \dots, X_{j_l}\} \subsetneq \{X_{i_1}, \dots, X_{i_k}\}} q(x_{j_1}, \dots, x_{j_l}) \\ &\quad \prod_{X_k \in \{X_{i_1}, \dots, X_{i_k}\} \setminus \{X_{j_1}, \dots, X_{j_l}\}} p(x_k). \end{aligned}$$

## Some notation

For example, we have

$$q(x_i) = 0,$$

$$q(x_i, x_j) = p(x_i, x_j) - p(x_i)p(x_j),$$

$$\begin{aligned} q(x_i, x_j, x_k) &= p(x_i, x_j, x_k) - p(x_i)p(x_j, x_k) - p(x_j)p(x_i, x_k) \\ &\quad - p(x_k)p(x_i, x_j) + 2p(x_i)p(x_j)p(x_k). \end{aligned}$$

With this notation, one can see that

$$q(x_{i_1}, \dots, x_{i_k}) = \sum_{t=1}^m w_t \prod_{j=1}^k A_{x_{i_j}}^t,$$

$$w^T A_{x_i} = 0,$$

where  $w = (w_1 \ \dots \ w_m)^T$  and  $A_{x_i} = (A_{x_i}^1 \ \dots \ A_{x_i}^m)^T$ .

$w$  is normal to the hyperplane spanned by the  $A_{x_i}$

Consider the parameters of a naive Bayesian distribution.  
Given vectors  $A_{u_1}, \dots, A_{u_{m-1}}$ , we have

$$(-1)^t \det(A_{u_1} \ \dots \ A_{u_{m-1}})^{\hat{t}} = w_t \det(1 \ A_{u_1} \ \dots \ A_{u_{m-1}}),$$

where the superscript  $\hat{t}$  denotes the removal of the  $t$ th row.

In other words, if

$$\det(1 \ A_{u_1} \ \dots \ A_{u_{m-1}}) \neq 0,$$

then  $w$  is the normal to the hyperplane spanned by  
 $A_{u_1}, \dots, A_{u_{m-1}}$  and whose components sum to 1.

The components of  $A_{x_i}$  are the roots of a degree  $m$  polynomial

For  $m = 2$ , we have

$$\begin{aligned} s^2 q(u_1, v_1) + sq(x_i, u_1, v_1) - q(x_i, u_1)q(x_i, v_1) \\ = q(u_1, v_1)(s + A_{x_i}^1)(s + A_{x_i}^2). \end{aligned}$$

For  $m = 3$ , we have

$$\begin{aligned} & s^3 \det \begin{pmatrix} q(u_1, v_1) & q(u_1, v_2) \\ q(u_2, v_1) & q(u_2, v_2) \end{pmatrix} + s^2 \left[ \det \begin{pmatrix} q(u_1, v_1) & q(u_1, v_2) \\ q(x_i, u_2, v_1) & q(x_i, u_2, v_2) \end{pmatrix} \right. \\ & \left. + \det \begin{pmatrix} q(x_i, u_1, v_1) & q(x_i, u_1, v_2) \\ q(u_2, v_1) & q(u_2, v_2) \end{pmatrix} \right] + s \left[ -\det \begin{pmatrix} q(u_1, v_1) & q(u_1, v_2) \\ q(x_i, u_2)q(x_i, v_1) & q(x_i, u_2)q(x_i, v_2) \end{pmatrix} \right. \\ & \left. - \det \begin{pmatrix} q(x_i, u_1)q(x_i, v_1) & q(x_i, u_1)q(x_i, v_2) \\ q(u_2, v_1) & q(u_2, v_2) \end{pmatrix} + \det \begin{pmatrix} q(x_i, u_1, v_1) & q(x_i, u_1, v_2) \\ q(x_i, u_2, v_1) & q(x_i, u_2, v_2) \end{pmatrix} \right] \\ & - \det \begin{pmatrix} q(x_i, u_1, v_1) & q(x_i, u_1, v_2) \\ q(x_i, u_2)q(x_i, v_1) & q(x_i, u_2)q(x_i, v_2) \end{pmatrix} - \det \begin{pmatrix} q(x_i, u_1)q(x_i, v_1) & q(x_i, u_1)q(x_i, v_2) \\ q(x_i, u_2, v_1) & q(x_i, u_2, v_2) \end{pmatrix} \\ & = \det \begin{pmatrix} q(u_1, v_1) & q(u_1, v_2) \\ q(u_2, v_1) & q(u_2, v_2) \end{pmatrix} (s + A_{x_i}^1)(s + A_{x_i}^2)(s + A_{x_i}^3). \end{aligned}$$

The components of  $A_{x_i}$  are the roots of a degree  $m$  polynomial

Given  $\mathbf{u} = \{u_1, \dots, u_{m-1}\}$  and  $\mathbf{v} = \{v_1, \dots, v_{m-1}\}$ , consider the polynomial of degree  $m$

$$\alpha_{x,\mathbf{u},\mathbf{v}}(s) = s^m \det \begin{pmatrix} q(u_1, v_1) & \cdots & q(u_1, v_{m-1}) \\ \vdots & & \vdots \\ q(u_{m-1}, v_1) & \cdots & q(u_{m-1}, v_{m-1}) \end{pmatrix} + s^{m-1} \cdots + s \cdots + \dots$$

whose coefficients are sums of determinants. We have

$$\alpha_{x,\mathbf{u},\mathbf{v}}(s) = \det \begin{pmatrix} q(u_1, v_1) & \cdots & q(u_1, v_{m-1}) \\ \vdots & & \vdots \\ q(u_{m-1}, v_1) & \cdots & q(u_{m-1}, v_{m-1}) \end{pmatrix} \prod_{t=1}^m (s + A_{x_i}^t).$$

## The parameters satisfy simple polynomial equations

Consider values  $\{x_1, \dots, x_k\}$ . The following equation holds

$$\det \begin{pmatrix} \prod_{j=1}^k A_{x_j}^t & q(x_1, \dots, x_k, v_1) & \dots & q(x_1, \dots, x_k, v_{m-1}) \\ A_{u_1}^t & q(u_1, v_1) & \dots & q(u_1, v_{m-1}) \\ \vdots & \vdots & & \vdots \\ A_{u_{m-1}}^t & q(u_{m-1}, v_1) & \dots & q(u_{m-1}, v_{m-1}) \end{pmatrix}$$
$$= q(x_1, \dots, x_k) \det \begin{pmatrix} q(u_1, v_1) & \dots & q(u_1, v_{m-1}) \\ \vdots & & \vdots \\ q(u_{m-1}, v_1) & \dots & q(u_{m-1}, v_{m-1}) \end{pmatrix}.$$

## The parameters satisfy simple polynomial equations

For  $\{x_1, \dots, x_k\} = \{u_0\}$ , we have

$$\det \begin{pmatrix} A_{u_0}^t & q(u_0, v_1) & \dots & q(u_0, v_{m-1}) \\ A_{u_1}^t & q(u_1, v_1) & \dots & q(u_1, v_{m-1}) \\ \vdots & \vdots & & \vdots \\ A_{u_{m-1}}^t & q(u_{m-1}, v_1) & \dots & q(u_{m-1}, v_{m-1}) \end{pmatrix} = 0.$$

For  $m = 3$  and  $\{x_1, \dots, x_k\} = \{u_1, u_2\}$ , we have

$$\begin{aligned} \det & \begin{pmatrix} A_{u_1}^t A_{u_2}^t & q(u_1, u_2, v_1) & q(u_1, u_2, v_2) \\ A_{u_1}^t & q(u_1, v_1) & q(u_1, v_2) \\ A_{u_2}^t & q(u_2, v_1) & q(u_2, v_2) \end{pmatrix} \\ &= q(u_1, u_2) \det \begin{pmatrix} q(u_1, v_1) & q(u_1, v_2) \\ q(u_2, v_1) & q(u_2, v_2) \end{pmatrix}. \end{aligned}$$

## Some determinants have an interpretable decomposition

Consider sets of values  $\mathbf{s}_1, \dots, \mathbf{s}_{m-1}$ . We have

$$\begin{aligned} \det & \begin{pmatrix} q(\mathbf{s}_1, v_1) & \dots & q(\mathbf{s}_1, v_{m-1}) \\ \vdots & & \vdots \\ q(\mathbf{s}_{m-1}, v_1) & \dots & q(\mathbf{s}_{m-1}, v_{m-1}) \end{pmatrix} \\ &= \left( \prod_{t=1}^m w_t \right) \det (1 \ A_{v_1} \ \dots \ A_{v_{m-1}}) \det M, \end{aligned}$$

where

$$M = \begin{pmatrix} 1 & \prod_{x \in \mathbf{s}_1} A_x^1 & \dots & \prod_{x \in \mathbf{s}_{m-1}} A_x^1 \\ \vdots & \vdots & & \vdots \\ 1 & \prod_{x \in \mathbf{s}_1} A_x^m & \dots & \prod_{x \in \mathbf{s}_{m-1}} A_x^m \end{pmatrix}$$

## Simple implicit equations follow

Consider a naive Bayesian distribution with  $m$  classes and consider sets of values  $\mathbf{s}_1, \dots, \mathbf{s}_{m'-1}$ . If  $m' > m$ , we have

$$\det \begin{pmatrix} q(\mathbf{s}_1, v_1) & \dots & q(\mathbf{s}_1, v_{m'-1}) \\ \vdots & & \vdots \\ q(\mathbf{s}_{m'-1}, v_1) & \dots & q(\mathbf{s}_{m'-1}, v_{m'-1}) \end{pmatrix} = 0.$$

Consider sets of values  $\mathbf{s}_1, \dots, \mathbf{s}_{m-1}$  and  $\mathbf{r}_1, \dots, \mathbf{r}_{m-1}$ . We have

$$\begin{aligned} & \det \begin{pmatrix} q(\mathbf{s}_1, v_1) & \dots & q(\mathbf{s}_1, v_{m-1}) \\ \vdots & & \vdots \\ q(\mathbf{s}_{m-1}, v_1) & \dots & q(\mathbf{s}_{m-1}, v_{m-1}) \end{pmatrix} \det \begin{pmatrix} q(\mathbf{r}_1, u_1) & \dots & q(\mathbf{r}_1, u_{m-1}) \\ \vdots & & \vdots \\ q(\mathbf{r}_{m-1}, u_1) & \dots & q(\mathbf{r}_{m-1}, u_{m-1}) \end{pmatrix} \\ &= \det \begin{pmatrix} q(\mathbf{s}_1, u_1) & \dots & q(\mathbf{s}_1, u_{m-1}) \\ \vdots & & \vdots \\ q(\mathbf{s}_{m-1}, u_1) & \dots & q(\mathbf{s}_{m-1}, u_{m-1}) \end{pmatrix} \det \begin{pmatrix} q(\mathbf{r}_1, v_1) & \dots & q(\mathbf{r}_1, v_{m-1}) \\ \vdots & & \vdots \\ q(\mathbf{r}_{m-1}, v_1) & \dots & q(\mathbf{r}_{m-1}, v_{m-1}) \end{pmatrix} \end{aligned}$$

# Outline

Mathematical Results

Applications

Extension

## Potential applications of our results

- Compute the set of parameters mapped to a given naive Bayesian distribution
- Estimate parameters from data by applying the previous computation to the distribution of observed frequencies
- Derive sufficient conditions for parameter identifiability and obtain results on the dimensionality of the model
- Building block in the computation of analytic asymptotic approximations to the marginal likelihood of the model
- Building block in model selection and learning of hidden causes

## An important hypothesis to compute the parameters

Suppose that we have a distribution  $p$  and sets of values

$$\mathbf{t} = \{t_1, \dots, t_{m-1}\},$$

$$\mathbf{u} = \{u_1, \dots, u_{m-1}\},$$

$$\mathbf{v} = \{v_1, \dots, v_{m-1}\}$$

such that

$$\det \begin{pmatrix} q(t_1, u_1) & \dots & q(t_1, u_{m-1}) \\ \vdots & & \vdots \\ q(t_{m-1}, u_1) & \dots & q(t_{m-1}, u_{m-1}) \end{pmatrix} \neq 0,$$

$$\det \begin{pmatrix} q(t_1, v_1) & \dots & q(t_1, v_{m-1}) \\ \vdots & & \vdots \\ q(t_{m-1}, v_1) & \dots & q(t_{m-1}, v_{m-1}) \end{pmatrix} \neq 0,$$

$$\det \begin{pmatrix} q(u_1, v_1) & \dots & q(u_1, v_{m-1}) \\ \vdots & & \vdots \\ q(u_{m-1}, v_1) & \dots & q(u_{m-1}, v_{m-1}) \end{pmatrix} \neq 0.$$

## Computation of $w$ from $A_{u_1}, \dots, A_{u_{m-1}}$

Our hypothesis amounts to

$$\begin{aligned} & \left( \prod_{i=1}^m w_i \right) \det \begin{pmatrix} 1 & A_{t_1} & \dots & A_{t_{m-1}} \end{pmatrix} \\ & \quad \det \begin{pmatrix} 1 & A_{u_1} & \dots & A_{u_{m-1}} \end{pmatrix} \\ & \quad \det \begin{pmatrix} 1 & A_{v_1} & \dots & A_{v_{m-1}} \end{pmatrix} \neq 0. \end{aligned}$$

Hence, we have

$$w_i = \frac{(-1)^i \det \begin{pmatrix} A_{u_1} & \dots & A_{u_{m-1}} \end{pmatrix}^{\hat{i}}}{\det \begin{pmatrix} 1 & A_{u_1} & \dots & A_{u_{m-1}} \end{pmatrix}}.$$

## Computation of $A_x$ from $A_{u_1}, \dots, A_{u_{m-1}}$

Since

$$\det \begin{pmatrix} A_{u_1}^t & q(u_1, v_1) & \dots & q(u_1, v_{m-1}) \\ \vdots & \vdots & & \vdots \\ A_{u_{m-1}}^t & q(u_{m-1}, v_1) & \dots & q(u_{m-1}, v_{m-1}) \\ A_x^t & q(x, v_1) & \dots & q(x, v_{m-1}) \end{pmatrix} = 0,$$

we have, for all values  $x$  distinct of  $v_1, \dots, v_{m-1}$ ,

$$A_x^T = (q(x, v_1) \ \dots \ q(x, v_{m-1})) \begin{pmatrix} q(u_1, v_1) & \dots & q(u_1, v_{m-1}) \\ \vdots & & \vdots \\ q(u_{m-1}, v_1) & \dots & q(u_{m-1}, v_{m-1}) \end{pmatrix}^{-1} (A_{u_1} \ \dots \ A_{u_{m-1}})^T.$$

## Computation of $A_{u_1}, \dots, A_{u_{m-1}}$

Find the roots of the polynomials  $\nu_{u_i, t, v}$  to obtain

$$\{A_{u_1}^1, \dots, A_{u_1}^m\},$$

⋮

$$\{A_{u_{m-1}}^1, \dots, A_{u_{m-1}}^m\}.$$

Note that these sets are not ordered: we are not able to assign each element to its hidden class.

There is some trivial non-identifiability due to the fact that classes can be permuted freely. To remove this degree of freedom from the analysis, we order the set  $\{A_{u_1}^1, \dots, A_{u_1}^m\}$  arbitrarily.

## Computation of $A_{u_1}, \dots, A_{u_{m-1}}$ : a brute force approach

For each ordering of each set  $\{A_{u_i}^1, \dots, A_{u_i}^m\}$  with  
 $i = 2, \dots, m - 1$

1. compute a candidate parameter with the previous procedure
2. test if the candidate satisfies the constraints to be a parameter and if it is mapped to the distribution

However, there are  $(m!)^{m-2}$  candidate parameters to test.

Corollary: under our hypothesis, there are at most  $(m!)^{m-1}$  parameters mapped to the distribution.

## Computation of $A_{u_1}, \dots, A_{u_{m-1}}$ : a second approach

We have

$$\begin{aligned} \det M \left( \sum_{p=1}^m \prod_{j=1}^k A_{x_{i_j}}^p \right) &= q(x_{i_1}, \dots, x_{i_k}) \det M \\ &+ \sum_{a=1}^{m-1} \sum_{b=1}^{m-1} (-1)^{a+b} q(x_{i_1}, \dots, x_{i_k}, t_a, v_b) \det M_{\hat{b}}^{\hat{a}}, \end{aligned}$$

where

$$M = \begin{pmatrix} q(t_1, v_1) & \dots & q(t_1, v_{m-1}) \\ \vdots & & \vdots \\ q(t_{m-1}, v_1) & \dots & q(t_{m-1}, v_{m-1}) \end{pmatrix} \quad (1)$$

We can constraint the orderings to those satisfying the above equation with  $\{x_{i_1}, \dots, x_{i_k}\} = \{u_1, u_j\}$ .

## Computation of $A_{u_1}, \dots, A_{u_{m-1}}$

The previous algorithm do not make use of all our theoretical results. For  $m = 3$ , recall that we have

$$\begin{aligned} \det \begin{pmatrix} A_{u_1}^t A_{u_2}^t & q(u_1, u_2, v_1) & q(u_1, u_2, v_2) \\ A_{u_1}^t & q(u_1, v_1) & q(u_1, v_2) \\ A_{u_2}^t & q(u_2, v_1) & q(u_2, v_2) \end{pmatrix} \\ = q(u_1, u_2) \det \begin{pmatrix} q(u_1, v_1) & q(u_1, v_2) \\ q(u_2, v_1) & q(u_2, v_2) \end{pmatrix}. \end{aligned}$$

We can derive  $A_{u_2}^t$  from  $A_{u_1}^t$  by solving the above equation.

We are currently investigating how to make use of all our results in the general case.

## The inversion algorithms can be adapted to estimate parameters

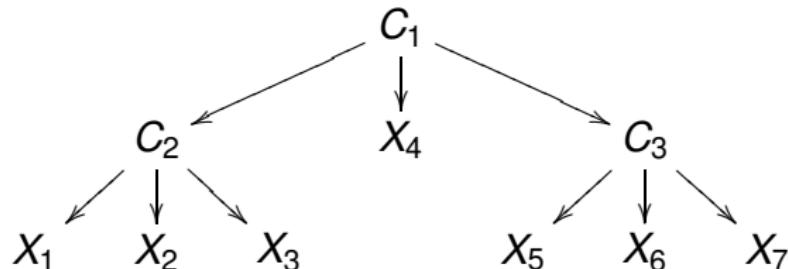
- Basic idea: apply the inversion algorithm to the observed distribution  $\hat{p}$ .
- Instead of testing whether a candidate parameter is mapped to  $p$ , we find the parameter minimizing the relative entropy to  $\hat{p}$ .
- Suppose that the unknown  $p$  is a naive Bayesian distribution with  $m$  classes satisfying our inversion assumption. As the sample size increases,  $\hat{p}$  converges to  $p$  and, by continuity, our estimate converges to a true parameter mapped to  $p$ .

## Practical issues

The estimation procedure has several issues:

- The computational complexity grows extremely fast with  $m$ , but linearly with  $n$ .
- The estimates are numerically unstable and require large sample sizes. For smaller sample sizes, there may not even be a single candidate parameter satisfying the parameter constraints.
- There are many degrees of freedom in the choice of  $\mathbf{t}$ ,  $\mathbf{u}$  and  $\mathbf{v}$ . Asymptotically, any choice is suitable. For small sample size, it is probably important.
- The results are not competitive with the E.M. algorithm.

## Extension to hierarchical latent class models



The parameters mapped to a HLC distribution with the above structure can be derived from the parameters mapped to the naive Bayesian distributions over

- $\{X_1, X_2, X_3\}$
- $\{X_1, X_4, X_5\}$
- $\{X_5, X_6, X_7\}$

obtained by marginalization.

## Conclusion

We presented some simple and interesting polynomial equations constraining a naive Bayesian distribution and its parameters. These results may be applied to

- compute the parameters mapped to a naive Bayesian distribution,
- estimate parameters from data.

The implicit equation

$$\det \begin{pmatrix} q(u_1, v_1) & \dots & q(u_1, v_{m'-1}) \\ \vdots & & \vdots \\ q(u_{m'-1}, v_1) & \dots & q(u_{m'-1}, v_{m'-1}) \end{pmatrix} = 0$$

holding for  $m' > m$  is similar to a tetrad constraint. A future research direction would investigate whether the constraint can indeed be used to learn hidden causes from data.