

# Geographical allozymes differentiation in wild *Phaseolus lunatus* L. of the Central Valley of Costa Rica and its implications for conservation and management of populations

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To suggest a conservation and management strategy for wild Lima bean (*Phaseolus lunatus* L.) in the Central Valley of Costa Rica, we examined the spatial distribution of genetic variation in 96 populations, using ten enzyme loci to analyse  $F$ -statistics and Moran's  $I$ . These loci displayed 20 alleles, of which 5 with relatively high frequencies were exclusively localised in the central part of the Valley. The estimates of  $F$ -statistics indicated a high level of genetic differentiation between populations (Mean  $F_{ST}=0.504\pm 0.094$ ). Such a value suggested that wild *P. lunatus* maintains about 50% of its genetic variation among populations. Moreover, the levels of inbreeding ( $F_{IT}=0.882\pm 0.026$  and  $F_{IS}=0.761\pm 0.012$ ) were high and significantly different from zero. Hence the genotypic composition of wild Lima bean deviated from Hardy-Weinberg proportions as a result of genetic differentiation between populations and non-random mating within populations. Spatial autocorrelation analysis using four loci showed positive and significant Moran's  $I$  at short distance in most cases. The resulting correlograms displayed up and down stochastic variations and indicated a patchy genetic structure. Combining the results obtained with those previously published on genetic structure, mating system, gene flow, and demography, we suggested probable causal factors and evolutionary mechanisms driving the genetic variability of the populations analysed. In addition, we indicated populations that should be preserved and proposed a reliable *in situ* management strategy.

**Keywords.** *Phaseolus lunatus*, Lima bean, autocorrelation, spatial genetic structure, *in situ* conservation.

**Différenciation géographique des allozymes chez la forme sauvage de *Phaseolus lunatus* L. dans la Vallée Centrale du Costa Rica et ses implications pour la conservation et la gestion des populations.** Afin de proposer une stratégie de conservation et de gestion pour la forme sauvage du haricot de Lima (*Phaseolus lunatus* L.) dans la Vallée Centrale du Costa Rica, nous avons examiné la distribution spatiale de la variabilité génétique de 96 populations. Dix loci enzymatiques ont été utilisés pour analyser les  $F$ -statistiques et l'indice  $I$  de Moran. Ces loci ont exprimé 20 allèles dont 5 ayant des fréquences relativement élevées étaient exclusivement localisés dans la partie centrale de la vallée. Les valeurs estimées des  $F$ -statistiques indiquaient un haut niveau de différenciation génétique entre les populations ( $F_{ST}=0.504\pm 0.094$ ). Une telle valeur suggère que la forme sauvage de *P. lunatus* maintient environ 50 % de sa variabilité totale entre les populations. En outre, les niveaux de consanguinité ( $F_{IT}=0.882\pm 0.026$  et  $F_{IS}=0.761\pm 0.012$ ) étaient significativement supérieurs à zéro. En conséquence, la composition génotypique du haricot de Lima sauvage dévie des proportions de Hardy-Weinberg suite à la différenciation génétique entre les populations et à l'absence de la panmixie dans les populations. L'analyse de l'autocorrelation spatiale effectuée à partir des données de quatre loci a montré que l'indice  $I$  de Moran était positif et significatif à courtes distances pour la plupart des loci. Les correlogrammes qui en résultent ont montré une variation stochastique irrégulière ainsi que l'existence d'une structuration spatiale de la diversité génétique en taches. En combinant les résultats des études précédentes concernant la structure génétique, le système de reproduction, le flux de gènes et la démographie, nous avons proposé des facteurs et mécanismes d'évolution qui pourraient moduler la variabilité génétique des populations analysées. En plus, nous avons indiqué les populations qui pourraient être protégées et proposé une stratégie fiable de gestion *in situ*.

**Mots-clés.** *Phaseolus lunatus*, haricot de Lima, autocorrelation, structure génétique spatiale, conservation *in situ*.

## 1. INTRODUCTION

In research on appropriate way to address plant resources conservation and management strategies, knowledge of the amount of genetic diversity and the spatial distribution of this diversity is essential for a correct diagnosis of the status, threats, and viability of populations (Hamrick, Allard, 1972; He *et al.*, 2000; Escudero *et al.*, 2003; Wilson, 2004). Genetic diversity may be spatially structured at different scales (geographic, population, subpopulation, etc.), due to environmental influence, life history, and demographic traits of the species (Loveless, Hamrick, 1984; Slatkin, 1985; Oostermeijer *et al.*, 2003). Consequently, spatial genetic structure provides a valuable tool for inferring causal factors and underlying operating evolutionary forces such as selection, gene flow, and drift (Nevo *et al.*, 1982; Nevo *et al.*, 1986; Barbujani, 1987; Epperson, 1990; Wilson, 2004).

Levin (1974), Loveless, Hamrick (1984), Wade, McCauley (1988), Oostermeijer *et al.* (1994) and Grassi *et al.* (2004) have highlighted the influence of environmental factors (including human activities and various interactions) and life history traits of plant species on population viability and genetic variability. An increasing number of studies have also integrated data from ecology, population biology, genetics, and reproductive biology in order to formulate reliable conservation and management strategies of populations (Schaal, Levin, 1976; Guerrant, 1992; Widén, 1993; Alvarez-Buylla *et al.*, 1996; Oostermeijer *et al.*, 2003). In these investigations, spatial analysis methods are of a special interest (Sokal, Oden, 1978a; Sokal, Oden, 1978b; Legendre, Fortin, 1989; Escudero *et al.*, 2003). Indeed, this technique may be helpful, for example, for the improvement of sampling strategies in collecting seeds for *ex situ* conservation, the selection of populations that should be protected *in situ*, the determination of a particular area size necessary for the conservation of a particular population, the selection of a specific site for the establishment of a corridor population, etc.

In 1992–2002, a wide research program describing population genetic structure, gene flow, genetic variability at geographical level, reproductive biology, and dynamics of wild Lima bean (*Phaseolus linaeus* L.) populations was conducted in the Central Valley of Costa Rica with the aim of developing a strategy for *in situ* conservation and management. This material represents a very important genetic reservoir for the improvement of the various *Phaseolus* bean cultigens, commonly found in many traditional cropping systems in Latin America and East Africa (Maquet, Baudoin, 1997). In this project, *P. linaeus* is also considered as a plant model due to its alternatively outbreeder-inbreeder behaviour. Indeed, Lima bean is a self-

compatible annual or short-living perennial species with a mixed-mating system, but predominantly self-pollinating, since the average outcrossing rate is low:  $0.096 \pm 0.071$  (Zoro Bi *et al.*, 2004).

Wild individuals in the valley are characterized by an indeterminate, climbing and vigorous growth habit, showing a prolonged flowering period (mid-November to mid-February) and a heavy pod load. Around 400 wild *P. linaeus* populations have been recorded in collaboration with the University of San José (Costa Rica) in the target area, which covers 2100 km<sup>2</sup>, in variants of premontane and lower montane humid forests, with altitudes ranging from 500 to 1800 m.a.s.l. Reproductive individuals can bear several racemes (about 400) with 1–20 pods per raceme, each pod containing 1–5 seeds. A three-year soil seed bank study indicated the occurrence of 3–5 seeds/m<sup>2</sup> according to populations (Degreef *et al.*, 2002). The annual germination rate from this soil seed bank ranged from 70 to 86%. In order to investigate the population genetic aspects, 22 enzyme loci from 15 enzymatic systems were resolved and their genetic basis established (Zoro Bi *et al.*, 1999). Sampling strategies integrating criteria of efficiency relevant to multilocus and many target populations have been investigated, in particular the number of plants and the number of seeds to be sampled (Zoro Bi *et al.*, 1998). Using the 22 enzyme loci resolved and the determined seeds sampling strategy, we analysed the genetic structure of 29 populations (Zoro Bi *et al.*, 2003). Thus, we quantified the proportion of polymorphic loci ( $P=10.32\%$ ), the mean number of alleles per locus ( $A=1.10$ ), and the mean effective number of alleles per locus ( $A_e=1.05$ ). The genotypic composition of the analysed populations showed deviation from the expected Hardy-Weinberg proportions. The total heterozygosity ( $H_T$ ), the intrapopulation genetic diversity ( $H_S$ ) and the interpopulation genetic diversity ( $D_{ST}$ ) were 0.193, 0.082, and 0.111 respectively. From the level of genetic differentiation between populations ( $F_{ST}=0.444$ ) which suggest that wild Lima bean maintains most of its isozyme variation among populations, gene flow was estimated, calculating the number of migrants per generation and assuming an island model (Wright, 1951):  $Nm=0.398$ .

In order to refine the diagnosis of genetic status and threats on wild Lima bean from the Central Valley of Costa Rica, additional investigations concerning the spatial structure of the genetic variation have been suggested. Here we report the results of a study aimed at determining the geographical distribution of allozyme frequencies, based on 96 wild Lima bean populations from the Central Valley of Costa Rica. Specifically, our goals were:

- to assess the patterns of the spatial distribution of the genetic diversity;

- to identify populations that were particularly interesting for conservation purposes;
- to suggest an *in situ* management strategy for the studied plant material.

## 2. MATERIAL AND METHODS

### 2.1. Plant material and sampling method

Ninety-six wild Lima bean populations distributed in the Central Valley of Costa Rica (Figure 1) were sampled between 1993 and 1998, from January to March, corresponding to the time of seed maturity. Each population was followed during a complete season so that all individuals bearing pods during this season were sampled. Geographical coordinates (Lambert projection) and associated vegetation of each sampled population were recorded (Table 1). The 29 populations analysed previously for the genetic structure study (Zoro Bi *et al.*, 2003) were included in the 96 selected populations. A population is here defined as any set of individuals, regardless of size that lives in the same habitat patch and isolated at least 500 m from other plants of the same species. We adopted a sampling strategy integrating criteria of efficiency relevant to a multilocus model and many target populations, designed for Lima bean (Zoro Bi *et al.*, 1998). Thus, in each selected population, we sampled all pod-bearing plants, resulting in sample sizes of one to 60 plants per population, and four to six racemes per plant. One seed was randomly chosen per raceme for electrophoretic analysis, resulting in sample sizes ranging from four to 334 seeds per population, according to the number of pod-bearing plants. The selected populations were identified by alpha-numeric codes.

### 2.2. Electrophoretic analysis

For electrophoretic variation, we analysed ten readable and reproducible enzyme loci resolved from eight enzymatic systems: alcohol dehydrogenase (ADH, E.C. 1.1.1.1), diaphorase (DIA, E.C. 1.8.1.4), esterases (fluorimetric and colorimetric: fEST and cEST, E.C. 3.1.1.-), glucose-6-phosphate isomerase (GPI, E.C. 5.3.1.9), malate dehydrogenase (MDH, E.C. 1.1.1.37), phosphogluconate dehydrogenase (PGDH, E.C. 1.1.1.44), phosphoglucomutase (PGM, E.C. 5.4.2.2), and shikimate dehydrogenase (SKDH, E.C. 1.1.1.25). Enzyme extraction was done by grinding 5-day-old cotyledon tissues in a potassium phosphate buffer, pH 7.0, containing 20% sucrose, 5% PVP-40, 0.05% triton X-100, 14 mM 2-mercaptoethanol, and 0.1 M  $\text{KH}_2\text{PO}_4$ . The pH value was adjusted to 7.0 with a solution of 5 M NaOH. Electrophoresis was performed using a horizontal 10% starch-gel containing 3% sucrose. Two buffer systems were employed: continuous histidine-citrate, pH 6.1 for ADH GPI, MDH, and PGDH, and discontinuous lithium-borate, pH 8.1/Tris-citrate, pH 8.4 for DIA, cEST, fEST, PGM, and SKDH. The techniques for gel electrophoresis and histochemical staining procedures are those reported elsewhere (Zoro Bi *et al.*, 1999).

Loci were labelled sequentially, with those migrating closest to the anodal end designated as number 1. Accession G25221 from the collection of the Centro Internacional de Agricultura Tropical (CIAT, Cali, Colombia), a Mexican wild form, was used as the control for our analyses. The allozyme from this genotype was designated 100 and all other allozymes were assessed according to their relative migration distance. The genetic control and the quaternary structure of the analysed enzyme systems have been discussed previously (Zoro Bi *et al.*, 1999).

### 2.3. Data analysis

The allozyme multilocus genotypes from polymorphic loci (99% criterion) were recorded and the obtained data were used to calculate allelic frequencies. We used *G*-tests (Sokal, Rohlf, 1995) to evaluate significant heterogeneity in allele frequencies among populations. *F*-statistics ( $F_{IT}$ ,  $F_{IS}$ , and  $F_{ST}$ ) estimated from genetic markers provide information on the genetic structuring within and among populations (Weir, Cockerham, 1984). Of these indices, the value of  $F_{ST}$  indicates how much of the genetic variation is partitioned among populations and then, can be used as a measure of the genetic differentiation that can be expected as a consequence of low level of gene flow among populations or steady differential selection (Slatkin, 1985; 1987). To analyse the genetic differentiation among the studied wild Lima bean populations,

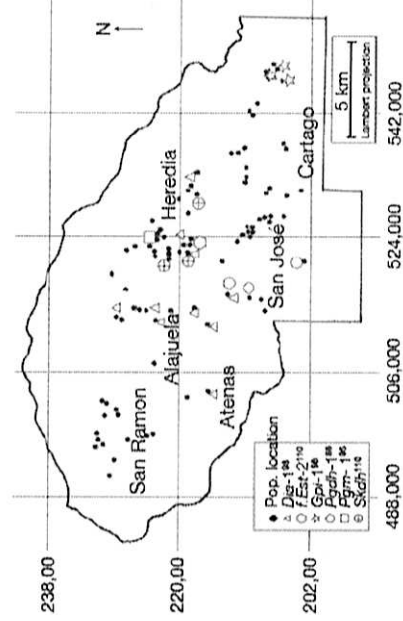


Figure 1. Spatial distribution of the less widespread alleles from six enzyme loci analysed in 96 wild Lima bean populations located in the Central Valley of Costa Rica — *Distribution spatiale des allèles rares de six loci enzymatiques analysés dans 96 populations sauvages du haricot de Lima localisées dans la Vallée Centrale du Costa Rica.*

**Table 1.** Geographic coordinates (Lambert projection), habitats, and estimated allele frequencies at 10 enzyme loci in 96 populations of wild *Phaseolus limatus* from the Central Valley of Costa Rica — *Coordonnées géographiques (projection Lambert), habitats, et fréquences alléliques estimées pour 10 loci enzymatiques sauvages de Phaseolus limatus provenant de la Vallée Centrale du Costa Rica.*

Pop.	N	Coordinates		Habitat	Frequency of allele controlling the fast migrating allozyme <sup>a</sup>											
		Latitude	Longitude		Adh-2 <sup>100</sup>	Dia-1 <sup>100</sup>	c.Est-2 <sup>100</sup>	f.Est-2 <sup>100</sup>	Gpi-1 <sup>100</sup>	Mdh-2 <sup>100</sup>	Pgdh-1 <sup>100</sup>	Pgm-1 <sup>100</sup>	Pgm-2 <sup>100</sup>	Skdh <sup>100</sup>		
A1	12	207.400	545.000	Bushy thicket	1	1	0	0	0	1	0	1	1	1	0	0
E19	5	208.500	522.300	Coffee plantation	1	1	0	0	0	1	1	1	1	1	0	0
E25	144	207.800	522.900	Bushy thicket	0.98	1	0	0	0	1	0.35	1	1	1	0.08	0
E27	5	210.500	521.300	Track	1	1	0	0	0	1	0.80	1	1	1	0	0
E28	5	211.700	512.200	Coffee plantation	1	0.30	0	0	0	1	0	0.20	1	1	0	0
E29	5	207.900	511.800	Bushy thicket	1	1	0	0	0	1	0.70	0.70	1	1	0	0
E30	5	205.800	509.500	Lining road	1	1	0	0	0	1	0	1	1	1	0	0
E31	15	206.500	525.200	Fallow	0.67	1	0	0	0	1	0	1	1	1	0.33	0
E34	5	206.300	525.100	Track	1	1	0	0	0	1	0	1	1	1	1	0
E35	16	205.200	525.600	Garden	1	1	-	0	0	1	1	1	1	1	0	0
E37	5	206.300	511.600	Bushy thicket	1	1	0	0	0	1	0	1	1	1	0	0
E50	86	205.100	524.700	Bushy thicket	0.82	1	0	0	0	1	0.30	1	1	1	0.23	0
E54	82	205.800	529.600	Garden	1	1	0	0	0	1	0.26	1	1	1	0.26	0
E56	5	204.900	530.300	Garden	1	1	0	0	0	1	1	1	1	1	0	0
E59	16	209.500	522.200	Garden	1	1	-	0	0	1	0	1	1	1	0	0
E65	5	218.200	521.200	Coffee plantation	0	1	0.10	0	0	1	0.10	1	1	1	0	0
E68	5	218.850	519.330	?	0.60	1	0.40	0	0	1	0	1	1	1	0.40	0
E76	58	219.200	520.300	Coffee plantation	0.89	0.73	0	0	0	1	0.10	1	1	1	0.78	0
E83	20	219.700	519.200	Coffee plantation	0.87	1	0	0	0	1	0.20	1	1	1	0	0
E84	109	218.500	519.100	Coffee plantation	0.43	0.93	0	0	0	1	0.78	1	1	1	0.21	0
E88	151	218.400	520.500	Track	0.43	1	0	0.01	0	1	0.50	1	1	1	0.60	0
E100	333	218.300	530.300	Coffee plantation	0.84	0.99	0	0	0	1	0.45	1	1	1	0.19	0
E104	5	202.000	528.900	Fallow	1	1	0	0	0	1	0	1	1	1	1	0
E110	16	204.800	550.900	Bushy thicket	0.19	1	0	0	0	1	0	1	1	1	0	0
E111	4	203.300	548.700	Lining road	1	1	0	0	0	1	0	1	1	1	0	0
E114	8	204.700	550.600	Bushy thicket	1	1	0	0	0	0	0	1	1	1	0	0
E115	5	218.400	494.700	Garden	1	1	0	0	0	1	0	1	1	1	0	0
G1	137	202.300	537.300	Coffee plantation	0.88	1	0	0	0	1	0.40	1	1	1	1	0
G7	5	202.500	537.600	Garden	1	1	0	0	0	1	0	1	1	1	0	0
G14	5	202.800	548.800	Track	1	1	0	0	0	1	0	1	1	1	0	0
G19	12	205.900	549.500	Lining road	1	1	0	0	0	1	1	1	1	1	0	0
G20	5	206.000	550.100	Track	0.80	1	0.20	0	0	0	0	1	1	1	0	0
G22	5	205.200	549.100	Bushy thicket	1	1	0	0	0	0	0	1	1	1	0	0
HER3	8	223.300	522.700	Coffee plantation	0.50	1	0.50	0	0	1	0.50	1	1	1	0	0
HER9	5	223.700	522.400	Coffee plantation	0	1	1	0	0	1	1	1	1	1	0	0
HER16	5	224.500	522.100	Lining road	0.50	1	0.40	0	0	1	0.40	1	1	1	0	0
HER30	5	224.700	524.400	?	0.70	1	0.20	0	0	1	0.80	1	0.60	0.60	0	0
HER45	5	223.800	521.100	Garden	0	1	0.60	0	0	1	0	1	1	1	0	0
J5	5	208.100	523.600	Garden	1	1	0	0	0	1	0.90	1	1	1	1	0
J6	10	208.400	522.800	Coffee plantation	0.90	1	0	0	0	1	0.70	1	1	1	0.55	0
J11	16	208.500	523.400	Garden	1	1	0	0	0	1	1	1	1	1	0.31	0
J29	218.400	529.900	-	0.10	1	0	0	0	0	1	0.90	1	1	0	0	0
J48	202	217.100	529.100	Coffee plantation	0.98	1	0	0	0	1	0.27	1	1	1	0.99	0.01
J58	5	230.100	483.900	Coffee plantation	1	1	0	0	0	1	0	1	1	1	1	0
J59	27	231.700	481.400	Coffee plantation	1	1	1	0	0	1	0	1	1	1	1	0
J67	5	226.700	509.400	Garden	0	1	0	0	0	1	0.90	1	1	1	0.20	0
J72	11	202.400	525.700	Garden	1	1	0.14	0	0	1	0.04	1	1	1	0	0
J87	102	199.500	530.000	Bushy thicket	1	1	0	0	0	1	0.20	1	1	1	0.01	0
KM12	12	199.100	517.900	Lining river	0.96	1	0.17	0	0	1	0.96	0.25	1	1	0	0
KM23	5	229.000	512.900	-	1	1	0	0	0	1	1	1	1	1	0	0
KM28	5	231.400	515.300	Bushy thicket	1	1	0	0	0	1	0	1	1	1	0	0
KM30	109	210.600	536.300	Bushy thicket	1	1	0	0	0	1	0.77	1	1	1	0.03	0.01
KM32	23	217.300	532.700	Coffee plantation	0	1	0	0	0	1	1	1	1	1	0	0
KM40	5	219.900	527.900	Coffee plantation	0.20	1	0	0	0	1	0.20	1	1	1	0.60	0

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Pop.	N	Coordinates	Habitat	Frequency of allele controlling the fast migrating allozyme <sup>1</sup>												
				Latitude	Longitude	Adh-2 <sup>100</sup>	Dia-1 <sup>100</sup>	c.Est-2 <sup>107</sup>	f.Est-2 <sup>119</sup>	Gpi-1 <sup>100</sup>	Mdh-2 <sup>120</sup>	Pgdh-1 <sup>100</sup>	Pgm-1 <sup>100</sup>	Pgm-2 <sup>100</sup>	Skdh <sup>130</sup>	
KM41	5	218.600	525.800	-	0	1	0	0	0	1	0	1	1	1	0	0
KM51	5	232.200	493.200	Coffee plantation	1	1	0	0	0	1	0	1	1	0	0	0
KM52	5	230.000	491.800	Coffee plantation	0.40	1	0.40	0	0	1	0	1	1	1	0	0
KM53	5	230.200	492.300	Coffee plantation	1	1	0.80	0	0	1	1	1	1	1	0	0
KM55	5	224.600	488.100	Lining road	1	1	1	0	0	1	0.60	1	1	1	0	0
KM56	5	224.200	488.100	Bushy thicket	0	1	0	0	0	1	0	1	1	1	0	0
KM57	5	228.400	487.300	-	1	1	1	0	0	1	0	1	1	0	0	0
KM60	5	223.600	500.300	-	0.20	1	0.20	0	0	1	1	1	1	1	0	0
KM62	5	223.800	500.200	Coffee plantation	1	1	0	0	0	1	0	1	1	1	0	0
KM63	31	232.700	493.500	Coffee plantation	1	1	0	0	0	1	1	1	1	1	0.02	0
KM67	5	221.200	509.200	Garden	1	1	0.20	0	0	1	1	1	1	1	0	0
P1	5	212.500	517.600	Lining road	1	1	0	0	0	1	1	1	1	1	0	0
P17	5	222.500	521.500	-	0	1	0	0	0	1	1	1	1	0	0	0
S10	5	223.400	507.300	Coffee plantation	0.90	0.80	0	0	0	1	1	1	1	1	0	0
S13	5	225.200	509.300	Lining river	1	0.20	0	0	0	1	0	1	1	1	0	0
S15	16	227.500	509.600	Garden	1	1	0.50	0	0	1	0	1	1	1	0	0
S17	5	230.200	508.000	-	1	1	1	0	0	1	0	1	1	1	0	0
S18	5	229.800	507.700	Garden	1	1	0.80	0	0	1	0	1	1	1	0	0
S21	5	221.900	519.500	Coffee plantation	0	1	0.75	0	0	1	0	1	1	1	0	0
S22	5	221.700	519.100	Coffee plantation	0	1	0.20	0	0	1	0	1	1	1	0	0.10
S23	10	221.800	518.500	-	0.90	0.60	0	0	0	1	0.90	1	1	1	0.50	0
S25	5	217.200	510.000	Lining road	1	1	0	0	0	1	1	1	1	1	0	0
S26	5	217.200	509.800	Lining road	1	1	0	0	0	1	1	1	1	1	0	0
S27	5	215.500	507.000	Lining road	1	0.40	0	0	0	1	1	1	1	1	0.40	0
S32	5	214.800	495.800	Lining road	1	0	0	0	0	1	1	1	1	1	0	0
SP19	5	209.300	531.500	Lining road	1	1	0	0	0	1	1	1	1	1	0	0
SP20	5	209.100	531.800	Lining road	1	1	0	0	0	1	1	1	1	1	0	0
SP21	5	209.200	533.800	Bushy thicket	1	1	1	0	0	1	1	1	1	1	0	0
SR8	5	233.800	486.000	Lining road	0	1	0	0	0	1	0	1	1	1	0	0
SR10	5	233.600	488.000	-	1	1	1	0	0	1	1	1	1	0	0	0
SR14	5	233.400	487.200	Garden	0	1	0	0	0	1	0	1	1	0	0	0
SR16	5	232.400	486.300	Lining road	1	1	0	0	0	1	0	1	1	1	0	0
SR20	5	232.600	491.000	Coffee plantation	0	1	0.20	0	0	1	0	1	1	1	0	0
ST5	5	228.900	519.400	Coffee plantation	0.60	0.20	1	0	0	1	0.20	1	1	1	0.40	0
ST7	5	227.700	520.200	Coffee plantation	1	1	0	0	0	1	0	1	1	1	0	0
ST11	5	225.000	520.100	Coffee plantation	1	1	0.50	0	0	1	0	1	1	1	0	0
ST14	5	224.100	518.300	Bushy thicket	0	1	0	0	0	1	1	1	1	1	0	0
ST44	12	224.200	518.700	Coffee plantation	0.71	1	0.33	0	0	1	0.96	1	1	1	0.33	0
TR23	5	209.400	537.100	Coffee plantation	1	1	0.25	0	0	1	1	1	1	1	0.80	0
TR36	5	211.800	536.200	Garden	1	1	0	0	0	1	0	1	1	1	0	0
TR54	137	208.700	543.600	Track	0.44	1	0	0	0	1	0.32	1	1	1	0.30	0
TR57	5	208.200	542.800	-	1	1	1	0	0	1	0.70	1	1	1	0	0

<sup>1</sup> Since all analysed loci were diallelic, only the frequency of the most anodally migrating allele is presented

<sup>2</sup> missing data

*F*-statistics were estimated for all polymorphic loci according to Weir and Cockerham (1984) and *F<sub>ST</sub>* was tested for significant difference from zero using a Student *t*-test (Sokal, Rohlf, 1995). Allele frequencies and *F*-statistics were computed using the GENEPOP software (Raymond, Rousset, 1995).

We tested whether geographical-scale spatial arrangement of allele frequencies was random using

spatial autocorrelation analysis (Cliff, Ord, 1973; Sokal, Oden, 1978a); results were presented as correlograms, i.e. graphics in which values of the autocorrelation coefficients are plotted against distance classes. Spatial autocorrelation analysis tests whether observations of a variable at one geographical site are independent of observations at neighbouring sites. The spatial autocorrelation was quantified using Moran's *I*

(1950) calculated from allele frequencies on individual populations. To perform the test, populations were connected according to Gabriel-connected scheme (Gabriel, Sokal, 1969). A total of 4560 connexions were then drawn. The performed Gabriel-connected graph was subdivided into 19 distance classes, each class corresponding to about 3 km. The distance classes number was chosen to ensure that at least 30 pairs of points were included for each class (Legendre, Fortin, 1989). Tests of significance (against the null hypothesis that allele frequencies were randomly distributed in space) were performed for each distance class by a randomisation process, whereas the overall significance of correlograms was tested using Bonferroni approximation. Calculation of Moran's  $I$  and the tests of significance were performed using the SAAP software (Wartenberg, 1989). Four enzyme loci (*Adh-2*, *c.Est-2*, *Mdh-2*, and *Pgm-2*) with two widespread alleles (both alleles observed in at least 15 populations) were selected to perform the spatial autocorrelation tests. For the other loci, expressing polymorphism in less than 15 populations (*Dia-1*, *f.Est-2*, *Gpi-1*, *Pgdh-1*, *Pgm-1*, and *Skdh*), we plotted the less common alleles locations on the sampling map to evaluate visually their spatial distribution.

### 3. RESULTS

#### 3.1. Populations' genetic differentiation

Each of the ten polymorphic loci analysed displayed two alleles, resulting in a total of 20 alleles observed. The most anodally migrating allozyme frequencies, calculated for each population and each locus are presented in table 1. The loci *f.Est-2* and *Pgm-1* expressed polymorphisms in only one population (E88 for *f.Est-2* and HER30 for *Pgm-1*), inducing skewed estimates for  $G$ -tests and inflated  $F$ -statistics

values. Consequently, these two loci were discarded in performing  $G$ -tests and  $F$ -statistics calculation.

Six out of eight loci tested showed significant allele frequencies heterogeneity among populations (Table 2). Non-significant  $G$  values were obtained with *c.Est-2* ( $G=52.92$ ,  $df=40$ ,  $P=0.065$ ) and *Pgdh-1* ( $G=8.69$ ,  $df=4$ ,  $P=0.101$ ). As expected on the basis of the previous studies (Zoro Bi *et al.*, 2003; 2004), the estimated  $F$ -statistics indicated a high level of genetic differentiation between populations (Mean  $F_{ST}=0.504\pm 0.094$ ), suggesting that wild Lima bean maintains about 50% of its isozyme variation among populations. Such tendency was confirmed by the fact that for all loci analysed except *Skdh*, the estimates of  $F_{ST}$  were significantly different from zero (Table 2), resulting in a high and significant mean value. Concomitantly,  $F_{IT}$  (the correlation between uniting gametes relative to all populations sampled) and  $F_{IS}$  (the correlation between uniting gametes within individual populations) were high and significant for the majority of the loci analysed and hence, for the mean values. Indeed, low values of  $F_{IT}$  and  $F_{IS}$  were observed only for *Skdh*, with negative sign for  $F_{IS}$ , suggesting an excess of heterozygotes at this locus. Contrary to *Skdh*, we obtained the highest estimates of  $F$ -statistics with *Gpi-1* ( $F_{IT}=F_{IS}=F_{ST}=1$ ), due to the fact that the two alleles observed at this locus were completely fixed in the sampled populations (i.e. in any population,  $p=1$  and  $q=0$  or  $p=0$  and  $q=1$ ). This suggested that the studied populations were completely differentiated at locus *Gpi-1*.

#### 3.2. Spatial distribution of allele frequencies

The less common alleles at six out of the ten loci analysed presented restricted geographic distribution. Indeed, besides the case of *f.Est-2* and *Pgm-1* explained before, the spatial distribution of the second allele from

Table 2. Tests of allele frequencies heterogeneity and  $F$ -statistics estimated for 96 populations of wild Lima bean — Tests d'hétérogénéité des fréquences alléliques et  $F$ -statistiques chez 96 populations sauvages du haricot de Lima.

Locus	G-test (df)	F-statistics $\pm$ SD		
		$F_{IT}$	$F_{IS}$	$F_{ST}$
<i>Adh-2</i>	529.24 (52)***	0.873 $\pm$ 0.026	0.778 $\pm$ 0.079	0.433 $\pm$ 0.079***
<i>Dia-1</i>	145.00 (16)***	0.921 $\pm$ 0.061	0.874 $\pm$ 0.111	0.376 $\pm$ 0.136*
<i>c.Est-2</i>	0.52.91 (40)	0.931 $\pm$ 0.040	0.744 $\pm$ 0.134	0.739 $\pm$ 0.098***
<i>Gpi-1</i>	026.61 (10)**	1.000 $\pm$ 0.000	1.000 $\pm$ 0.000	1.000 $\pm$ 0.007***
<i>Mdh-2</i>	402.16 (62)***	0.823 $\pm$ 0.035	0.747 $\pm$ 0.038	0.299 $\pm$ 0.077***
<i>Pgdh-1</i>	008.69 (4)	0.997 $\pm$ 0.194	0.815 $\pm$ 0.565	0.741 $\pm$ 0.075**
<i>Pgm-2</i>	874.11 (48)***	0.917 $\pm$ 0.035	0.777 $\pm$ 0.065	0.625 $\pm$ 0.122***
<i>Skdh</i>	012.73 (4)*	0.001 $\pm$ 0.002	-0.013 $\pm$ 0.040	0.014 $\pm$ 0.042
Mean <sup>1</sup>		0.882 $\pm$ 0.026	0.761 $\pm$ 0.012	0.504 $\pm$ 0.094***

<sup>1</sup> Standard deviations (SD) of locus data. The grand mean as well as the mean's SD were calculated using numerical resampling (jackknife) over populations (locus data) or over loci (mean and its SD). \*  $P < 0.05$ . \*\*  $P < 0.01$ . and \*\*\*  $P < 0.001$ . The comparison is based on G-tests for allelic frequencies among populations and Student t-tests for FST.

the four other loci (*Dia*-1, *Gpi*-1, *Pgdh*-1, and *Skdh*) showed an interesting particularity for conservation purpose (Figure 1): *Dia*-1<sup>98</sup>, *Pgdh*-1<sup>98</sup>, and *Skdh*<sup>110</sup> were exclusively located in the central part, and *Gpi*-1<sup>98</sup> in the western part of the target site. Dry season in the western part is not as severe as in the central part.

Spatial autocorrelation analysis was performed for alleles from four enzyme loci: *Adh*-2, *c.Est*-2, *Mdh*-2, and *Pgm*-2. The resulting correlograms are presented in figure 2. Positive autocorrelation was found at shorter distance for the main loci analysed and this was significant for *Adh*-2 (3-9 km) and *c.Est*-2 (3-6 km). Then, the autocorrelation decreased to become negative at distance of 9-39 km, with significant values for *Adh*-2 (12-30 km). At large distances (50-60 km), positive autocorrelation was observed again for alleles from *Mdh*-2 and *Pgm*-2; for the later, significant value was obtained at 54 km. For *Adh*-2 and *c.Est*-2, autocorrelation remained negative at the same lag classes, the trend being significant with *c.Est*-2 (51-54 km). Overall, the correlograms for the four individual loci showed an irregular shape, displaying up and down stochastic variations. Two correlograms, namely those of *Adh*-2 and *c.Est*-2, were significant (Figure 2). The pattern of correlograms indicated a patchy structure of allozyme variants in wild *P. lunatus* from the Central Valley of Costa Rica.

The results show that isolation by distance occurs in these populations.

#### 4. DISCUSSION

Of the 20 alleles recorded, six were sporadic and localised, with two private alleles (as defined by Slatkin, 1985): *f.Est*-2<sup>110</sup> and *Pgm*-1<sup>95</sup>. Two aspects of the results concerning these alleles appeared particularly interesting for conservation. First, data given in table 1 showed that except *f.Est*-2<sup>110</sup> and *Skdh*<sup>110</sup> (frequencies=0.01 and 0.01-0.10, respectively), all the alleles with a restricted distribution presented relatively high frequencies in most populations in which they were observed: 0.01-0.80 for *Dia*-1<sup>98</sup>, 1 for *Gpi*-1<sup>98</sup>, 0.30-0.80 for *Pgdh*-1<sup>98</sup>, and 0.40 for *Pgm*-1<sup>95</sup>. The second aspect of these results was the grouping of five of these alleles (*Dia*-1<sup>98</sup>, *f.Est*-2<sup>110</sup>, *Pgdh*-1<sup>98</sup>, *Pgm*-1<sup>95</sup>, and *Skdh*<sup>110</sup>) in the central part of the Valley, *Gpi*-1<sup>98</sup> being localised in the western part, characterized by a less severe dry season compared to the central part (Figure 1). These results could be related with the breeding system of the wild Lima bean populations. Indeed such pattern of genetic variation is expected in general for any predominantly self-pollinating species (because alleles often tend to be fixed within population) displaying isolation by distance (because rare alleles will usually be restricted to a reduced geographical

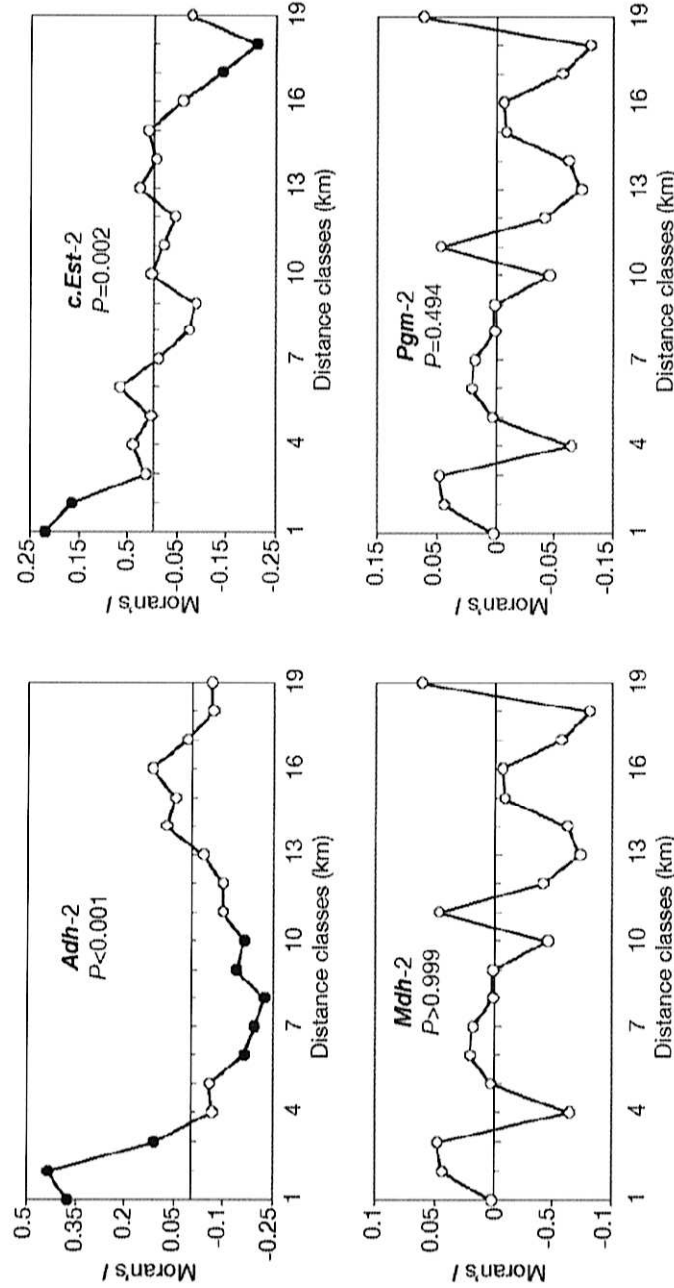


Figure 2. Correlograms of Moran's *I* for four enzyme loci analysed in 96 wild Lima bean populations from the Central Valley of Costa Rica. Filled circles represent Moran's *I* that are significantly different from zero ( $\alpha = 0.05$ ). *P* is the overall significance of the correlogram — Correlogrammes de *I* de Moran pour quatre loci enzymatiques analysés dans 96 populations sauvages du haricot de Lima de la Vallée Centrale de Costa Rica. Les cercles remplis représentent les *I* de Moran qui sont significativement différents de zéro ( $\alpha = 0.05$ ). *P* est la signification globale du correlogramme.

range due to limited dispersal). Another hypothesis to explain the two aspects of our results is the impact of selection. Genotypes with localised alleles could have relatively high viability or reproductive success, leading to high frequencies of these alleles in the populations involved. It has been established from studies of natural selection using various predominantly self-pollinated plant species that homozygote genotypes at a marker locus may differ substantially in fitness according to life cycle stages, seasons, and environment (Allard, Workman, 1963; Imam, Allard, 1965; Hamrick, Allard, 1972; Nevo *et al.*, 1982; Nevo *et al.*, 1986; Ennos, 1990). Simulation studies (Epperson, 1990) have also shown that selection can greatly contribute to the development of the spatial genetic structuring in plant populations. For wild Lima bean, the restricted geographic distribution of the sporadic alleles could indicate the occurrence of microhabitat-level alleles selection in the populations studied: as a result, genotypes carrying locally advantaged alleles might have high survival value, due to natural selection tending to increase the homozygosity. This hypothesis could be supported by data on *Gpi-1<sup>96</sup>* which, although rare (observed in only three populations), was completely fixed (frequency=1). As will be discussed more thoroughly later, random genetic drift in small bottlenecked populations and founder effects resulting from extinction/recolonisation episodes that characterised the populations studied could enhance the fixation process of the advantaged alleles. Indeed, these populations are usually found in open and disturbed areas with grasses and scattered trees or bushy thickets; they also colonise coffee plantations from the long-living fences (usually *Erythrina* and euphorbs) bordering the plots. Each year, some wild Lima bean populations are eliminated due to land management, such as growing urbanisation, severe grazing, seasonal fires in pasture lands and sugar cane plantations, and the replacement of traditional small-scale coffee plantations by modern high input demanding plantations (Rocha *et al.*, 1997). Recolonisation of the cleared sites could be due to nearby plants or to new individuals emerging from the soil seed bank.

Since the correlation between uniting gametes relative to all populations as expressed by  $F_{IT}$  was high and significant ( $F_{IT}=0.882\pm 0.026$ ), we deduced that the genotypic composition of wild Lima bean showed a deviation from the Hardy-Weinberg proportions. This disequilibrium was attributable to both genetic differentiation between populations ( $F_{ST}=0.504\pm 0.094$ ) and non-random mating within populations ( $F_{IS}=0.761\pm 0.012$ ). The review of Hamrick and Godt (1990) based on eight life histories and ecological traits of plants, and presenting a compilation of statistics on populations' genetic structure indicates that in short-lived perennial and predominantly selfing species, gene

differentiation among populations is high. Lima bean is a mixed-mating and predominantly autogamous species (Zoro Bi *et al.*, 2004) that is expected to express a high level of genetic divergence among populations, coupled with an important heterozygote deficiency. The estimates of the populations' genetic structure indices analysed in our study were in accordance with the designated trend. It should also be noted that the estimated  $F$ -statistics were close to those obtained from previous study using 29 populations and nine polymorphic loci (Zoro Bi *et al.*, 2003):  $F_{IT}=0.932\pm 0.066$ ,  $F_{IS}=0.866\pm 0.128$ , and  $F_{ST}=0.497\pm 0.358$ . Factors explaining the genetic structuring in wild Lima bean populations and hypotheses on the evolutionary processes likely to affect them have been thoroughly discussed in the indicated paper. Briefly, it has been argued that the most likely phenomena explaining the low allelic richness, the frequent and steady heterozygotes deficiency, and the high genetic divergence of wild Lima bean populations are founder effects, genetic drift, high selfing rate, Wahlund effects, and limited gene flow between populations.

Study of Rocha *et al.* (1997) has highlighted genetic drift and founder effects in the studied populations. Indeed, the authors established after a 7-year survey that of the 400 populations inventoried in the target site, about 60 (only 16%) contained more than five pod-bearing plants, so that genetic drift must be high. The presence of the founder effects is supported by the fact that during the period between 1992 and 1998, the number of plants reaching reproductive age differs markedly among years, varying from one to 50 plants in one population (data not shown).

In the Central Valley of Costa Rica, the outcrossing rate of wild Lima bean is low, ranging from 0.027 to 0.268, with a mean of  $0.096\pm 0.071$  (Zoro Bi *et al.*, 2004). Such estimates suggested that this plant had a high level of autogamy.

The Wahlund effects, i.e. the existence of genetic structure in a population, were mentioned as another factor explaining the frequent deficiency of heterozygotes observed. The actual genetic structure of wild Lima bean was assessed in a previous work using isozyme electrophoresis and three populations (Zoro Bi *et al.*, 1997). Seeds were sampled according to a grid of  $4\times 4$  m for bidimensional populations or 4 m apart for linear populations and the genotypes of mother plants at each node were so determined. The genetic structure in the populations was obvious: alternative alleles at each locus were clustered in opposite parts of the populations, creating a patch structure mainly composed of homozygote individuals (Wahlund effects).

From the estimates of  $F_{ST}$  ( $=0.504$ ) and on the basis of Wright's (1951) equation, as modified by Crow and Aoki (1984), the number of migrants per generation was  $Nm=0.243$ . This value must however be taken with



caution because Wright's equation is established for an island model at equilibrium and our populations are unlikely to be at equilibrium.

Since most of the genetic variation in wild Lima bean proved to reside among populations, the question of how it is spatially organised appeared relevant. *F*-statistics provide information on the genetic structuring within and among populations. Tests of hypotheses on the evolutionary processes that account for the observed gene frequencies distribution, however, require more complex approaches such as spatial autocorrelation (Sokal, Oden, 1978a; Barbujani, 1987; Epperson, 1990; Heywood, 1991). The significant spatial autocorrelation of allele frequencies at shorter distance classes observed for most of the loci analysed indicated clearly that these variables were not spatially random in wild Lima bean from the Central Valley of Costa Rica. Recent reviews on genetic structure studies in plant populations (Sokal, Oden, 1978a; Barbujani, 1987; Heywood, 1991; Escudero *et al.*, 2003; Vekemans, Hardy, 2004) have addressed differences in spatial distribution of genetic variation at various life stages or age classes and in different microenvironments of a particular population or zone. For wild Lima bean, the single-locus genetic structure due to isolation by distance is highly stochastic and may well explain alone variation among loci. Other mechanisms influencing the genetic structuring of the studied populations cannot be discarded, such as isolation-by-distance due to limited pollen and seed flow between populations, founder effects, genetic drift, and microhabitat selection of alleles. In our case, the dissimilar course of correlograms curves could result from combined effects of the factors suggested, creating various patch structures in wild Lima bean.

## 5. IMPLICATIONS FOR CONSERVATION AND MANAGEMENT

Knowledge of the genetic variability within a taxon is crucial for conservation purposes, when interpreted within a broader ecological and organismal context. For wild *P. limatus* populations, isozyme electrophoresis data indicated high genetic heterogeneity among populations, localised alleles, and patchy structures of allele frequencies throughout the sampled zone. The occurrence of patchy genetic structure show the relevance of *in situ* conservation actions for these populations. The results also suggest that many populations distributed throughout the range of the species should be protected, since the conservation of few populations would not guarantee the preservation of a representative sample of the existing genetic diversity, as alleles distribution greatly varied between patches. Thus, protecting the following populations

should preserve the less common and localised alleles identified: E28, E29, E76, E84, E88, E114, G22, HER30, J48, KM12, S10, S13, S23, S27 and ST5. Moreover, if populations E35, G1, G19, J59 and KM32 are included, the preserved allelic variation would be increased to 100%.

Once populations have been selected for *in situ* conservation, their sound management is necessary to preserve a high level of genetic variability. For wild Lima bean populations, an appropriate management method was indicated from demographic and soil seed banks studies (Degreef *et al.*, 1997; Degreef *et al.*, 2002). These authors reported that most wild *P. limatus* populations from the target zone produced adult individuals every year and were characterised by abundant seeds production. However, rainfall during the post-ripening period or adult plants destruction by man or cattle could markedly reduce seeds production and then, the number of individuals in such populations the next year. Seed dormancy and soil seed banks in these populations ensured the recolonisation of the sites. Based on information gained from these studies, it was suggested to pay particular attention to the timing of clearing and weeding which are regularly carried out in the sites. Indeed, weeding coffee plantations in which wild Lima bean populations are localised just after the seed dispersal (at the end of the dry season) favours the breakdown of seed dormancy and reduces germination delays, because these practices expose the seeds to high temperatures and humidity at the moment of the first rains. Clearing was proposed at the beginning of the rain season, since field observations showed that populations cut or disturbed during the dry season were less likely to regrow and accordingly, less likely to produce seeds in the next fruiting period.

The most reliable way for the preservation of wild *P. limatus* genetic variability in the Central Valley of Costa Rica is the design of synthetic populations with all the allelic diversity identified, given that natural populations are threatened, due to several human activities in this zone. In 1998, synthetic populations have been established in protected sites throughout the Valley. These synthetic populations should contain genotypes carrying all the alleles identified. Preliminary demographic study on the synthetic populations confirmed that a careful management is required to break seed dormancy. Weeding during rainy season could speed up the colonisation process and ensure the stability of these populations (Meurrens *et al.*, 2001).

It must be recognised that allozymes often underestimate levels of intra- and interpopulations genetic variation for adaptive traits crucial to the survival and reproduction of plants (Hamrick *et al.*, 1991; Francisco-Ortega *et al.*, 2000). Consequently, further investigations using more variable genetic

markers such as microsatellites, AFLPs, RADPs or ISSRs (Schaal *et al.*, 1991; Amos, Hoelzel, 1992, Ouedraogo *et al.*, 2005) are required to refine the present results and suggestions for the conservation and management of wild *P. linnatus* and other plant species with similar biological and ecological traits.

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