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A cosmology-independent calibration of Type Ia supernovae data

C. Hauret, * P. Magain and J. Biernaux

STAR-OrCA, Université de Liège, Allée du 6 Août 19c, B-4000 Liège, Belgium

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ABSTRACT

Recently, the common methodology used to transform Type Ia supernovae (SNe Ia) into genuine standard candles has been suffering criticism. Indeed, it assumes a particular cosmological model [namely the flat Λ cold dark matter (Λ CDM)] to calibrate the standardization corrections parameters, i.e. the dependence of the supernova peak absolute magnitude on its colour, post-maximum decline rate and host galaxy mass. As a result, this assumption could make the data compliant to the assumed cosmology and thus nullify all works previously conducted on model comparison. In this work, we verify the viability of these hypotheses by developing a cosmology-independent approach to standardize SNe Ia data from the recent JLA compilation. Our resulting corrections turn out to be very close to the ΛCDM-based corrections. Therefore, even if a ACDM-based calibration is questionable from a theoretical point of view, the potential compliance of SNe Ia data does not happen in practice for the JLA compilation. Previous works of model comparison based on these data do not have to be called into question. However, as this cosmology-independent standardization method has the same degree of complexity than the model-dependent one, it is worth using it in future works, especially if smaller samples are considered, such as the superluminous Type Ic supernovae.

Key words: methods: data analysis – supernovae: general – cosmology: miscellaneous.

1 INTRODUCTION

Since the end of the 1990s, modern cosmology has been developed around the idea that the most accurate representation of the contents and evolution of our Universe is the flat Λ cold dark matter (Λ CDM) model, a model of a spatially flat universe containing baryonic and dark matter as well as a repulsive component called dark energy. Over the years, this model proved to be extremely predictive and consistent with the many and varied observations of Type Ia supernovae (SNe Ia), cosmological microwave background or large-scale structures, to name but a few. Nowadays, this model is so well established that it has become the standard model of cosmology, totally outshining any other cosmological model.

At the origin of the 20-yr reign of this model, we find the observations of SNe Ia (Riess et al. 1998; Perlmutter et al. 1999). These objects are extremely important for cosmologists as they are nearly perfect standard candles, observable over large distances. An astrophysical object is a standard candle if its intrinsic luminosity is known or can be determined independently of its distance. One can thus deduce the distance of such an object by measuring its apparent flux.

SNe Ia are believed to arise from thermonuclear explosions of white dwarfs in a binary system (Hoyle & Fowler 1960). As these white dwarfs accrete matter from their companions, they grow and

reach explosion conditions when their mass approaches the Chandrasekhar limit (Chandrasekhar 1931). If the very details of this scenario are still a matter of debate, it is reasonable to assume that all SNe Ia present roughly the same luminosity as the explosions of white dwarfs of similar masses should produce the same amount of energy. Therefore, these objects make good standard candles.

However, in the last decades, small but significant variations of their peak luminosities have been observed, implying that some corrections have to be applied in order to use SNe Ia as genuine standard candles. Correlations have been found between the intrinsic brightness of SNe Ia, the shape of their light curve (Phillips 1993), their colour (Tripp 1998), and their host galaxy (Kelly et al. 2010; Lampeitl et al. 2010; Sullivan et al. 2010). Thanks to these correlations, SNe Ia are nowadays considered as one of the best cosmological tools.

The methodology to deal with these light-curve shape, colour, and host galaxy corrections (called hereafter the standardization corrections) is extremely important in order to accurately describe our Universe via cosmological models. Over the years, a particular approach appeared (Perlmutter et al. 1999) and developed amongst the community. Nowadays widely approved and used (e.g. Suzuki et al. 2012; Betoule et al. 2014; Calcino & Davis 2017, for recent examples), this methodology is unfortunately not perfect. Indeed, it assumes a cosmological model, having the potential collateral effect of making the data compliant to this particular model, disfavouring any other cosmologies (Gopal Vishwakarma & Narlikar 2010; Mar-

^{*} E-mail: clementine.hauret@ulg.ac.be

riner et al. 2011; Melia 2012; Wei et al. 2015a). Corrected that way, the SNe Ia could no longer be used to compare different models of the Universe and works to determine the most representative one (such as the ones recently published by Xu & Zhang 2016; Hee et al. 2017) could then be void and obsolete.

In this paper, we develop a simple and completely cosmology-independent approach to standardize SNe Ia data and to check whether this problem actually occurs in the JLA compilation data¹ (Betoule et al. 2014). Hence in Section 2, we explain in further details the currently used methodology to process SNe Ia standardization corrections and we highlight its potential issues. Then, we present our alternative calibration and answer the question of the possible compliance of data in Section 3.

2 THE CURRENT METHODOLOGY TO PROCESS SNe Ia DATA

In order to be considered as genuine standard candles, the SNe Ia peak luminosities have to be corrected from the standardization corrections. Mathematically speaking, following the works of Phillips (1993), Tripp (1998), and Suzuki et al. (2012), we can compute the corrected peak absolute magnitude as

$$M_{\rm B,corr} = M_{\rm B} - \alpha x_1 + \beta c + \delta P \left(M_{\rm stellar} < 10^{10} \,\mathrm{M}_{\odot} \right). \tag{1}$$

 $M_{\rm B}$ is the absolute blue magnitude of a Type Ia supernova (SN Ia) with x_1 , c, and P=0 while $M_{\rm B,\,corr}$ is this magnitude for an SN Ia of given x_1 , c, and P. The latter are parameters linked respectively to the SN Ia light-curve shape, its colour, and its host galaxy mass (see below), while α , β , and δ are parameters that describe the correlations of the peak magnitude to the three aforementioned properties.

First, x_1 is a measurement of the light-curve shape, related to the so-called stretch correction as the differences in light-curve shape can also be seen as the stretching of the light-curve time axis (Perlmutter et al. 1997a,b). Secondly, c is a parameter characterizing the colour of the object (Tripp 1998). Finally, P is the probability that the SN Ia host galaxy is less massive than a threshold fixed at $10^{10} \,\mathrm{M}_{\odot}$ (Sullivan et al. 2010). One will notice that the relation between host mass and absolute magnitude presented in equation (1) can be implemented differently than with the probability P defined in this paper or in others such as Suzuki et al. (2012) or Mohlabeng & Ralston (2014). For example, it can be computed as a step function of the host mass as it has been done in Betoule et al. (2014) or in the very recent Calcino & Davis (2017).

Initially, $M_{\rm B}$ as well as the α and β coefficients² were calibrated on nearby SNe Ia (e.g. Phillips 1993; Hamuy et al. 1995; Tripp 1997, 1998, as well as papers based on the MLCS light-curve fitter introduced by Riess, Press & Kirshner 1996) and the relations were applied to more distant objects. However, since Perlmutter et al. (1999), another determination of these standardization corrections has been introduced and developed around the SALT and SALT2 light-curve fitters (Guy et al. 2005, 2007). Nowadays, this methodology is the most common way to transform SNe Ia into standard candles. $M_{\rm B}$, α , β , and δ are seen as nuisance parameters and are determined together with the cosmological parameters by fitting an assumed model (generally the flat Λ CDM model) on the Hubble

diagram (Suzuki et al. 2012; Betoule et al. 2014; Calcino & Davis 2017).

Unfortunately, while it gives better fits if the flat ΛCDM model is in fact the best representation of our Universe, this technique can lead to a bias in its favour (Gopal Vishwakarma & Narlikar 2010; Marriner et al. 2011; Melia 2012; Wei et al. 2015a). Indeed, when the cosmology and the standardization corrections are fitted simultaneously, their respective parameters are no longer independent from each other. In theory, this could skew the observational data towards the assumed cosmology (a flat Λ CDM model in the case we are interested in), at the expense of any other one. This will especially be true if the standardization corrections are correlated with the redshift. With such corrections, the SNe Ia data could no longer be used to test cosmological models as they would always favour the assumed one. In order to correctly compare the ability of various cosmological models to fit the SNe Ia data, the standardization parameters should be redetermined for each non-ΛCDM cosmologies (as it has been done by Wei et al. 2015a but not in model comparison works we are interested in such as Xu & Zhang 2016; Hee et al. 2017), leading to possible time-consuming calculations.

Of course, nowadays the flat Λ CDM model is widely favoured by various observations. Hence, most of the present observational works, including SNe Ia analyses, are not aimed any more at testing which cosmological model is the most representative of our Universe but they are instead designed to increase the precision on the determination of its density parameter $\Omega_{m,\,0}$. In this case, the methodology currently used for SNe Ia is perfectly well suited. However, its possible collateral effect on data and on model comparison raises important philosophical questions. Should we approve of a way of processing the data that can potentially make them compliant to a certain theoretical model? Should we not prefer a cosmology-independent approach to avoid this possible problematic effect?

3 DEVELOPING A COSMOLOGY-INDEPENDENT CALIBRATION OF SNe Ia DATA

To answer these questions and check if this potential compliance of data is effectively present in the SNe Ia compilations, we develop an alternative cosmology-independent calibration to process the standardization corrections. Different cosmological models could then be fitted on the recalibrated SNe Ia data in order to determine the most representative model of our Universe as it has been done by Xu & Zhang (2016) for example. The decisive advantage of the standardization parameters values we determine is that they can be directly used when working with non-ΛCDM cosmologies instead of having to be re-evaluated for each studied cosmological model.

3.1 Calibration principles

Apart from developing a precise model able to determine the standardization parameters directly from the physics of the SNe Ia thermonuclear explosions, the only way to evaluate these parameters is through the Hubble diagram. Indeed, on the one hand, one can express the distance modulus μ of each SN Ia as a difference between its apparent and absolute magnitude (respectively m_B and $M_{\rm B, corr}$) and, using equation (1), we have

$$\mu = m_{\rm B} - M_{\rm B} + \alpha x_1 - \beta c - \delta P \left(M_{\rm stellar} < 10^{10} \,\rm M_{\odot} \right).$$
 (2)

¹http://supernovae.in2p3.fr/sdss_snls_jla/ReadMe.html

²One can notice that the δ coefficient was not yet introduced at that time but it can nevertheless be calibrated the same way.

On the other hand, this distance modulus can also be found from the SN Ia redshift z via the luminosity distance $d_L(z)$:

$$\mu = 5\log(d_{\rm L}(z)) - 5.$$
 (3)

Unfortunately, the $d_L(z)$ relation cannot be found exactly without assuming a particular cosmological model. Hence, as previously said, the recent SNe Ia compilations have been primarily processed assuming a flat Λ CDM model but one can use other cosmology, such as the $R_h = ct$ model (Wei et al. 2015a) for example. However, this prior choice of a model could lead to the already-pointed-out problem of data compliance.

To put this possibility to the test, we decide to get rid of any assumption on the cosmology at the cost of losing the exact dependence of luminosity distance with redshift. To do so, we expand the $d_L(z)$ relation as a Taylor series around the observer (for whom z = 0). We test two different Taylor expansions: (i) one of the redshift (hereafter called the *z*-series):

$$d_{L}(z) = d_{0,z} \left(z + d_{1,z} z^{2} + d_{2,z} z^{3} + d_{3,z} z^{4} + O(z^{5}) \right)$$
(4)

and (ii) one of a function f of the redshift (f(z) = z/1 + z) (hereafter called the f-series):

$$d_{L}(z) = d_{0,f} \left(f(z) + d_{1,f} f^{2}(z) + d_{2,f} f^{3}(z) + d_{3,f} f^{4}(z) + O\left(f^{5}(z)\right) \right).$$
 (5)

One can directly notice that in both cases, we logically impose a zero luminosity distance at a zero redshift. In Section 3.3, we will choose between these two kinds of series, limited to different orders, so as to minimize the number of Taylor parameters while optimizing the representation of SNe Ia data.

Our specific choice of the f(z) function can be explained as follows. As SNe Ia with redshifts larger than unity have been observed, the high-order terms of the z-series could quickly become non-negligible if their associated parameters $\{d_{i,z}\}$ are non-null. The $d_L(z)$ relation would then not be well approximated by a limited series and our calibration would be useless. This is why we also consider the f-series with the f function such that it stays low as z goes high. We would then a priori neglect high-order terms more easily than in the z-series. Furthermore, one can notice that in a model based on the Friedmann–Lemaître–Robertson–Walker (FLRW) metric, this f-series is in fact the luminosity distance expanded as a Taylor series of the scale factor R of the Universe around its actual value $R(t_0)$ as $z + 1 = R(t_0)/R$. Our particular choice of the function f(z) is thus quite natural in an FLRW context.

To sum up, with these two Taylor series, all the information about cosmology is exclusively included in the Taylor parameters $\{d_{i,z/f}\}$. We would like to point out that in neither of our series, do we suppose a specific form of the metric. We do not even assume General Relativity. Nonetheless, for cosmological models based on General Relativity and on the FLRW metric, these parameters are directly related to the well-known present Hubble H_0 and deceleration q_0 parameters defined by

$$H_0 \equiv H(t_0) = \frac{\dot{R}(t_0)}{R(t_0)}$$
 and $q_0 \equiv q(t_0) = -\frac{\ddot{R}(t_0)}{R(t_0)} \left[\frac{\dot{R}(t_0)}{R(t_0)}\right]^{-2}$.

Hereafter, we will set the scale factor present value $R(t_0)$ at unity. In this way, the first two Taylor parameters of the *z*-series have a very simple form:

$$d_{0,z} = \frac{c_{\text{light}}}{H_0}$$
 and $d_{1,z} = \frac{1 - q_0}{2}$,

where c_{light} is the speed of light. For completeness, the expressions of the next two Taylor parameters of this same series can be found in Visser (2004).

Going back to our estimate of the standardization parameters through the Hubble diagram, we can combine equations (2), (3), (4) and (5) (limiting here the series to the third order) to have

$$m_{\rm B} - M_{\rm B} + \alpha x_1 - \beta c - \delta P \left(M_{\rm stellar} < 10^{10} \,\mathrm{M}_{\odot} \right)$$

 $\simeq 5 \log(d_{0,z}) + 5 \log \left(z + d_{1,z} z^2 + d_{2,z} z^3 \right) - 5$ (6)

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$$m_{\rm B} - M_{\rm B} + \alpha x_1 - \beta c - \delta P \left(M_{\rm stellar} < 10^{10} \,\mathrm{M}_{\odot} \right)$$

 $\simeq 5 \log(d_{0,f}) + 5 \log \left(f(z) + d_{1,f} f^2(z) + d_{2,f} f^3(z) \right) - 5. (7)$

The terms appearing in these equations are divided in two categories: (i) the SN Ia observable properties $(m_B, x_1, c, P, \text{ and } z)$ and (ii) the standardization and Taylor parameters $(\alpha, \beta, \delta, d_{i \neq 0, zlf}$ as well as a combination³ of M_B and $d_{0, zlf}$). The latter can be determined by minimizing a χ^2 statistics (as it has been made in Suzuki et al. 2012; Betoule et al. 2014) or, in a more rigorous approach, by using a Bayesian framework (such as the one first developed by March et al. 2011 and its extensions applied to the most recent SNe Ia compilation by Shariff et al. 2016, for example).

3.2 Data set and intrinsic dispersion

In this work, we use SNe Ia data released in the JLA compilation from Betoule et al. (2014), containing 740 objects with redshifts ranging from around 0.01 to 1.3.

When working with SNe Ia data, one has to evaluate a parameter besides the standardization parameters: the intrinsic dispersion of the SNe Ia magnitudes. It may include both a true intrinsic scattering component and any uncharacterized sources of errors.

To determine its value, we cannot use the usual methodology (for examples Conley et al. 2011; Sullivan et al. 2011; Suzuki et al. 2012), which consists in adding an error to that of the SNe Ia apparent magnitude until the best-fitting χ^2 reaches a value of one per degree of freedom. This procedure obviously forbids any comparison between models as all fits end up with equally good (artificial) χ^2 values. This problem is not new and has been pointed out by many authors such as Conley et al. (2011). Nevertheless, it is particularly an issue for our work as we eventually have to choose between the different Taylor series presented in the previous subsection.

Fortunately, Betoule et al. (2014) developed a technique to avoid this problem and to evaluate this dispersion more properly. Basically, they divide the SNe Ia into different redshift bins and evaluate the intrinsic dispersion in each bin via a restricted log-likelihood criterion. For more information, we strongly encourage the reader to directly turn to section 5.5 of Betoule et al. (2014).

Our resulting intrinsic dispersion can be found in Table 1. One will notice that similarly to Betoule et al. (2014) and authors before them (Conley et al. 2011; Sullivan et al. 2011), we determine different dispersions for each sample of objects at our disposal [i.e. the low-z, Sloane Digital Sky Survey-II (SDSS-II), SNLS (Supernova Legacy Survey), and *Hubble Space Telescope (HST)* samples from the JLA compilation]. Indeed, the dispersion we want to evaluate includes a true intrinsic component as well as a component coming from the remaining unexplained sources of errors, which could vary

 $^{^3}$ For models based on the FLRW metric, we find the well-known degeneracy between $M_{\rm B}$ and $H_{\rm 0}$.

Table 1. Magnitude intrinsic dispersion evaluated separately on each SNe Ia sample through the methodology developed by Betoule et al. (2014). Our values are compatible with previous determinations (Conley et al. 2011; Betoule et al. 2014).

Sample	Intrinsic dispersion
low-z	0.116
SDSS-II	0.045
SNLS	0.128
HST	0.060

from one survey program to another. Then, the dispersion associated to each sample in Table 1 is the weighted mean of the dispersions evaluated on the different redshift bins defined by Betoule et al. (2014).

3.3 Cosmology-independent calibration

As stated in the two previous subsections, we have to choose between the $d_{\rm L}$ z- and f-series (equations 4 and 5) and limit them to a particular order, so that we minimize the number of Taylor parameters while optimizing the representation of our data.

We calibrate the standardization and Taylor parameters on SNe Ia data for these two series limited to five increasing orders. To determine the optimal of these 10 calibrations, we do not use the usual minimization of a χ^2 statistics as this criterion does not take into account the complexity (i.e. the number of parameters) of the series. Indeed, at equivalent agreement, we want to favour a model with fewer parameters. This has also the advantage to avoid the

overtraining of our methodology. Other statistical criteria, called information criteria, have been developed to consider this idea of complexity. Among them, the Akaike information criterion (AIC; Akaike 1974) and the Bayesian information criterion (BIC; Schwarz 1978) are widely used. They are defined by

$$AIC = -2 \ln \mathcal{L}_{max} + 2k$$
 and $BIC = -2 \ln \mathcal{L}_{max} + k \ln N$,

where \mathcal{L}_{max} is the maximum likelihood ($-2 \ln \mathcal{L}_{\text{max}} = \chi^2_{\text{min}}$ if we assume that the errors on the data are normally distributed, as it is generally the case in cosmology), k is the number of parameters of the studied model, and N is the number of data points.

The model with the lowest AIC and BIC will be preferred but only the relative values between different models are important and useful for comparison. Thus, to quantify the goodness of a given model i compared to the one with the minimum AIC or BIC, one has to compute the differences $\Delta \text{AIC}_i = \text{AIC}_i - \text{AIC}_{\min}$ or $\Delta \text{BIC}_i = \text{BIC}_i - \text{BIC}_{\min}$. The larger these differences, the stronger the empirical evidence against the given model. More information about these criteria can be found in Liddle (2004) for example.

Our resulting standardization and Taylor parameters can be found in Table 2. Each set of parameters is determined using a simulated annealing algorithm implementing a χ^2 minimization on the whole SNe Ia population, i.e. on the 740 objects. For each of these sets, the covariance matrix has been calculated, yielding the one-sigma error bars. Table 3 displays the corresponding χ^2 for these best-fitting models and their associated AIC and BIC, so that models can be compared to each other. These $d_{\rm L}(z)$ series are also plotted on Hubble diagrams, compared to binned data, on Fig. 1.

Table 2. Best-fitting standardization and Taylor parameters for our cosmology-independent methodology, up to different orders of both z and f Taylor series. In bold, one can find the calibration minimizing the number of Taylor parameters while optimizing the data representation (see Table 3 associated).

	$M_{ m B}+{ m 5log}(d_0,z_{lf})$										
	Order	α	β	δ	- 5	$d_{1,z/f}$	$d_{2, z/f}$	$d_{3, z/f}$	$d_{4, z/f}$		
z-series	1st	0.121 ± 0.007	4.58 ± 0.11	0.177 ± 0.015	24.300 ± 0.008	_	_	_	_		
	2nd	0.149 ± 0.006	3.358 ± 0.071	0.085 ± 0.012	24.074 ± 0.009	0.606 ± 0.018	_	_	_		
	3rd	0.149 ± 0.006	$\boldsymbol{3.369 \pm 0.071}$	$\boldsymbol{0.084 \pm 0.012}$	24.049 ± 0.012	$\boldsymbol{0.726 \pm 0.044}$	-0.151 ± 0.050	_	_		
	4th	0.150 ± 0.006	3.370 ± 0.071	0.084 ± 0.012	24.044 ± 0.015	0.761 ± 0.075	-0.26 ± 0.19	0.08 ± 0.13	_		
	5th	0.150 ± 0.006	3.372 ± 0.071	0.084 ± 0.012	24.041 ± 0.018	0.79 ± 0.13	-0.41 ± 0.55	0.33 ± 0.84	-0.12 ± 0.39		
f-series	1st	-0.381 ± 0.065	33.2 ± 2.8	0.68 ± 0.10	24.934 ± 0.063	_	_	_	_		
	2nd	0.158 ± 0.006	3.924 ± 0.084	0.098 ± 0.013	23.822 ± 0.013	3.500 ± 0.060	_	_	_		
	3rd	0.148 ± 0.006	3.393 ± 0.071	0.087 ± 0.012	24.093 ± 0.016	1.118 ± 0.089	5.49 ± 0.22	_	_		
	4th	0.150 ± 0.006	3.368 ± 0.071	0.084 ± 0.012	24.036 ± 0.021	1.96 ± 0.22	0.3 ± 1.3	8.7 ± 2.0	_		
	5th	0.150 ± 0.006	3.369 ± 0.071	0.084 ± 0.012	24.038 ± 0.027	1.88 ± 0.42	1.4 ± 3.8	3 ± 13	7 ± 14		

Table 3. χ^2 , χ^2 per degree of freedom (dof), and information criteria (AIC and BIC) associated to the best-fitting parameters values of Table 2. For each of these calibrations, the fits have been made on all the N = 740 SNe Ia. In bold, one can find the model minimizing the number of Taylor parameters while optimizing the data representation (minimum values of both AIC and BIC).

	Order	k	χ^2	χ^2/dof	AIC	ΔAIC	BIC	$\Delta \mathrm{BIC}$
z-series	1st	4	2555	3.47	2563	1522	2581	1512
	2nd	5	1038	1.41	1048	7	1071	2
	3rd	6	1029	1.40	1041	0	1069	0
	4th	7	1029	1.40	1043	2	1075	6
	5th	8	1029	1.41	1045	4	1082	13
f-series	1st	4	4385	5.96	4393	3352	4411	3342
	2nd	5	1556	2.21	1566	525	1589	520
	3rd	6	1055	1.44	1067	26	1095	26
	4th	7	1031	1.41	1045	4	1077	8
	5th	8	1030	1.41	1046	5	1083	14

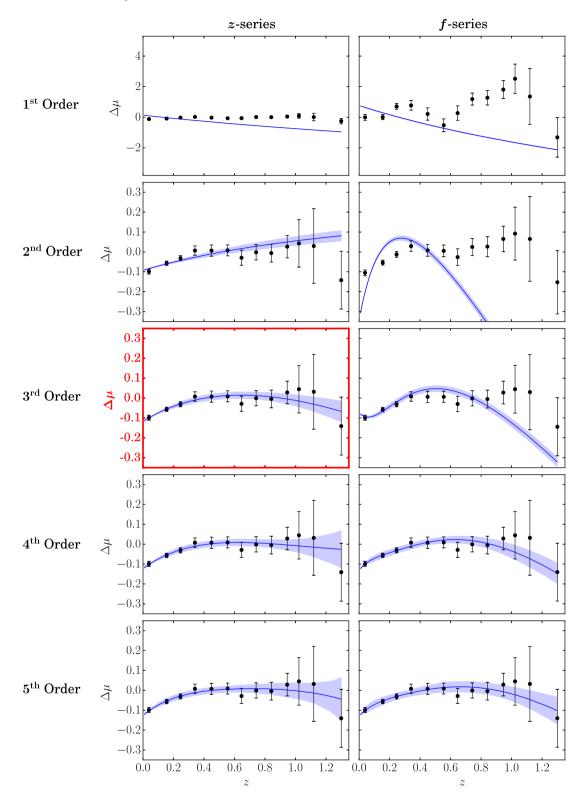


Figure 1. Hubble diagram residuals when the distance moduli of an empty cosmological model have been subtracted. Each panel compares the weighted SNe Ia data over 0.1-length redshift bins (built on the standardization parameters presented in Table 2; black points) with our calibration distance moduli (built on the Taylor parameters presented in this same table; blue line with shaded zones corresponding to 1σ errors). From left to right: calibration based on the *z*-or *f*-series. From top to bottom: calibration based on the first to the fifth order of the series. Obviously, the *f*-series reproduce the behaviour of the data less correctly than the *z*-ones. Amongst the latter, the series limited to the third order (whose panel is highlighted in bold and red) minimizes the number of Taylor parameters while optimizing the representation of the data (see Table 3).

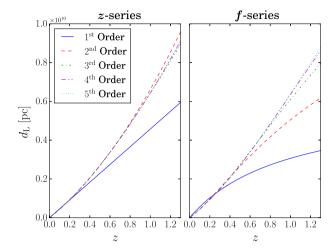


Figure 2. Evolution of our best-fitting z- (left-hand panel) and f-series (right-hand panel) limited to different orders as a function of the redshift z. These approximated $d_L(z)$ relations are given by equations (4) and (5) whose Taylor parameters values $\{d_{i,zlf}\}$ can be found in Table 2. The addition of supplementary terms in the z-series results in smaller modifications of the $d_L(z)$ function than for the f-series. Despite an a priori preference for the f-series to expand the $d_L(z)$ function, the z one eventually appears to be the one to use in this situation.

Table 4. Covariance matrix associated to our best cosmology-independent calibration (based on the z Taylor series limited to the third order). The standardization and Taylor parameters are presented in the same order than in Table 2.

$$\mathrm{Cov} = 10^{-6} \times \begin{pmatrix} 32 & -60 & 21 & -13 & 22 & -16 \\ & 5056 & 34 & -45 & 425 & -303 \\ & & 147 & -46 & -32 & 9 \\ & & & 147 & -436 & 410 \\ & & & & 1929 & -1984 \\ & & & & & 2457 \end{pmatrix}$$

First of all, it is clear in Table 3 and Fig. 1 that the *z*-series represent the data more quickly and more correctly than the *f*-ones. Even the fifth order of the *f*-series does not provide a fit as good as the *z*-series, third order. This result is slightly surprising as one would *a priori* be more confident in the *f*-series where the values f(z) = z/1 + z stays low while *z* itself can reach relatively high values (there are nine objects with $z \ge 1$ in our data set). Non-negligible differences between the true $d_L(z)$ relation and its Taylor approximation in *z* could thus be expected at high redshift. Nevertheless, if the redshift powers of equation (4) can be high for the most distant objects, these powers are multiplied by relatively low Taylor parameters $\{d_{i,z}\}$ as can be seen in Table 2. This has to be compared with the lower powers of f(z) (see equation 5), multiplied by the quite high associated $\{d_{i,f}\}$ values, presented in this same table.

This can also be visualized in Fig. 2 where the z and f Taylor series limited to different orders as a function of the redshift are plotted. One can directly see that adding terms in the z-series has a weaker effect on the approximated $d_L(z)$ function than in the f-one. Hence, even though the z-series seems at first more likely to invalidate the Taylor approximation, it turns out to give better results than the f-ones.

Consequently, having decided which kind of series to use in our calibration, its order is left to be selected. As can be seen in Table 3, limiting the z-series to the third order gives optimal values of both the information criteria, AIC and BIC. Furthermore, in Table 2, the $d_{3,z}$ and $d_{4,z}$ Taylor parameters associated to the fourth- and fifth-order series are not significantly different from zero.

On top of that, one can notice, when comparing lines of Table 2 or when observing the (unmoving) binned data points on Fig. 1, that the standardization parameters α , β , and δ do not significantly vary starting from the second order of the *z*-series or from the third one of the *f*-series. This is also illustrated on Fig. 3, showing the evolution of these parameters with the order of the Taylor series. Their values quickly stabilize for both the series. We could then even use the standardization parameters values given by the fit of the second order of the *z*-series. One may notice that even if our calibration had suffered from overtraining, the parameters we are interested in (i.e. the standardization parameters) would not have been impacted by this issue.

For purposes of rigour, our cosmology-independent standardization parameters will be chosen as the ones found through our *z*-series limited to the third order (highlighted in bold in our Tables 2 and 3). The covariances are given in the matrix presented in Table 4. The SNe Ia data can then be properly recalibrated and used to perform fits of various cosmological models. Xu & Zhang (2016), for example, developed such a work on the raw JLA data. We want to highlight the fact that if using the same SNe Ia data set to perform the calibration and the cosmological fits, the error propagation between the parameters of these two fits is quite tricky. We thus strongly encourage the reader to separate the SNe Ia at their disposal into two independent subsets, performing the calibration fit on one group of objects and the cosmological one on the other.

3.4 Digression about cosmology-dependent calibrations

For the sake of illustration, we also performed cosmology-dependent fits on SNe Ia data. These calibrations simultaneously fit the standardization parameters as well as the cosmological ones as it was done by Suzuki et al. (2012), Betoule et al. (2014), and many others since Perlmutter et al. (1999). These parameters as well as their associated χ^2 , AIC, and BIC values are displayed in Table 5, along with ΔAIC_{CI} and ΔBIC_{CI} , the differences between the AIC/BIC values of the particular cosmology-dependent fit and those values for our best cosmology-independent (CI) calibration. From these values, one can notice that the cosmology-dependent calibrations have a goodness-of-fit similar to the best independent ones, the flat Λ CDM fit being marginally favoured by the data due to its smallest number of parameters.

With this paper, we aim at testing the possible compliance of the SNe Ia data when fitting standardization and cosmological parameters simultaneously. We thus develop a cosmology-independent calibration to compare results from dependent and independent methods. For the sake of this comparison, a clear distinction is made between those two families of calibrations: when focusing on finding the best independent one, we favour our Taylor expansion, even though the cosmology-dependent calibration results in slightly better fits (as it is the case here for the flat Λ CDM model). On the other hand, when using SNe Ia data in general cosmological tests, as done by Shi, Huang & Lu (2012) or Xu & Zhang (2016) for example, and in case of a cosmology-dependent calibration leading to a much better fit than an independent one, we would advise to favour the former, despite the philosophical questions raised.

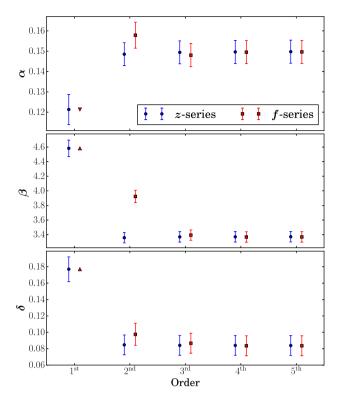


Figure 3. Evolution of the standardization parameters (from top to bottom: α , β , and δ) as a function of the order of the limited z- (blue circles) and f-series (red squares). Exact values can be found in Table 2. For purposes of readability, the parameters of the f-series limited to the first order are not depicted on the panels. Red triangles indicate their out-of-graph location. One can notice the non-significant variation of the parameters values from the second order of the z-series and from the third order of the f-ones.

3.5 Discussion

For readability, the standardization parameters that we found via our best cosmology-independent calibration as well as via a simultaneous fit with the flat Λ CDM model are summarized in Table 6. Parameters from Betoule et al. (2014) are also presented.

Comparing the first two lines of Table 6, one can see that we retrieve quasi-identical parameters using either the simultaneous fit methodology or our cosmology-independent one. This means that even though assuming a particular cosmology to standardize the SNe Ia could theoretically lead to the data compliance problem, it is not the case in practice for the JLA compilation. JLA data are not significantly skewed towards the flat Λ CDM model assumed to process them. We thus independently confirm the results from Betoule et al. (2014), where the possible correlation between the cosmological and the standardization parameters was found to be non-existent for this compilation.

This can be explained by the fact that the numerous SNe Ia from the JLA compilation sufficiently populate the space of standardization parameters and redshift. However, if that conclusion is valid for the present SNe Ia compilation, which includes hundreds of objects, it might not be valid for other standard candles, for which the sample might not be as large and as independent on redshift. As an example, let us mention the superluminous supernovae (Type Ic), for which only about 20 are presently known (Inserra & Smartt 2014, 2015; Wei, Wu & Melia 2015b). In such smaller samples, there is no guarantee that the space of standardization parameters and redshift will be sufficiently filled so that the calibration will be independent of the assumed cosmological model.

4 CONCLUSION

Even though the cosmology-dependence of SNe Ia data does not effectively appear in the JLA compilation, the philosophical interrogations raised at the end of Section 2 still remain open, valid, and unsettled. In a sounder and more rigorous theoretical perspective, we would thus recommend the use of our simple, general, and cosmology-independent methodology to determine the standardization parameters, particularly because it can be applied to any standard candle sample. Our method could then be used in future works comparing cosmological model predictions to standard candle results, whatever the nature and size of the sample.

Table 5. Best-fitting standardization and cosmological parameters along with their associated χ^2 , χ^2 per degree of freedom (dof), AIC, and BIC for simultaneous fits of different basic cosmologies. The present Hubble constant is assumed to be equal to $70 \, \mathrm{km \, s^{-1} \, Mpc^{-1}}$. The $\Delta \mathrm{AIC_{CI}}$ (respectively $\Delta \mathrm{BIC_{CI}}$) are defined as the differences between the corresponding AIC (resp. BIC) associated to the studied cosmology and the AIC (resp. BIC) of the best cosmology-independent (CI) calibration (here, the one built on the z-series limited to the third order). These latter can be found in bold in Table 3.

Cosmology	α	β	δ	$M_{ m B}$	$\Omega_{\mathrm{m,0}}$	$\Omega_{\Lambda,0}$	w	k	χ^2	χ ² /dof	AIC	ΔAIC_{CI}	BIC	ΔBIC _{CI}
Flat ACDM	0.150 ± 0.006	3.379 ± 0.071	0.084 ± 0.012	-19.130 ± 0.008	0.273 ± 0.018	-	-	5	1030	1.40	1040	-1	1063	-6
Λ CDM	0.149 ± 0.006	3.361 ± 0.071	0.084 ± 0.012	-19.108 ± 0.013	0.095 ± 0.095	0.44 ± 0.14	_	6	1030	1.40	1042	+1	1070	+1
Flat wCDM	0.149 ± 0.005	3.363 ± 0.061	0.084 ± 0.012	-19.108 ± 0.012	$0.01^{+0.17}_{-0.01}$	_	-0.60 ± 0.15	6	1030	1.40	1042	+1	1070	+1
wCDM	0.150 ± 0.005	3.371 ± 0.061	0.084 ± 0.012	-19.124 ± 0.017	0.15 ± 0.21	0.36 ± 0.57	-1.5 ± 2.1	7	1029	1.40	1043	+2	1075	+6

Table 6. Best-fitting standardization parameters for our cosmology-independent methodology (copied from the line highlighted in bold in Table 2) and from a simultaneous fit technique similar to the one used by Betoule et al. (2014) (copied from the first line of Table 5). The M_B parameter values are given assuming a present Hubble constant of 70 km s⁻¹Mpc⁻¹. Our parameter values are also compared to values from Betoule et al. (2014) analysis. One will notice some significant differences.

	Methodology	α	β	δ	$M_{ m B}$
Present work	Cosmo-indpt	0.149 ± 0.006	3.369 ± 0.077	0.084 ± 0.012	-19.11 ± 0.01
	Sim. fit (flat ΛCDM)	0.150 ± 0.006	3.379 ± 0.071	0.084 ± 0.012	-19.13 ± 0.01
Betoule et al. (2014)	Sim. fit (flat Λ CDM)	0.141 ± 0.006	3.101 ± 0.075	0.070 ± 0.023	-19.05 ± 0.02

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